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Is It What You Inherited Or What You Learnt?

Intergenerational linkage and interpersonal
inequality in Senegal

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Abstract

Using original survey data on Senegal that include an individualized measure of consumption, we study the role played by land inheritance, other bequests and parental background as influences on an adult's economic welfare and economic activities. While intergenerational linkages are evident, we find a seemingly high degree of mobility across generations, associated with the shift from farm to non-farm sectors and the greater economic activity of women. Male-dominated bequests of land and housing bring little gain to mean consumption and play little role in explaining inequality. Inheritance of non-land assets and the education and occupation of parents (especially the mother) and their choices about children's schooling are more important to adult welfare than property inheritance. Significant gender inequality in consumption is evident.

Keywords: inheritance, land, mobility, inequality, gender
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Tables appear at the end of the paper.

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1. Introduction

The reallocation of productive assets to more efficient uses is an important means of promoting higher aggregate output in any economy. The most important (non-labour) asset in most poor countries is land. The economic, social and political means by which a country allocates its land can thus have a bearing on aggregate economic efficiency. Much of the literature has argued that well-defined individual land rights are crucial for efficiency, though there have been dissenting views.¹

The main means of changing land ownership rights studied in the literature are market purchases and administrative land reallocations, both of which have been studied extensively.² Yet, in many developing-country settings it may well be the case that land inheritance is a more common means of re-assigning land-use rights, although this has received very little attention. With imperfect credit markets, inheritance may have a bearing on the efficiency of farming and also promote new non-farm economic activity by providing the kind of lumpy boost to individual wealth that can allow new investment opportunities to be pursued—opportunities that would otherwise be impossible to take up given borrowing constraints. Thus there is interest in better understanding inheritance.

Equity is another reason we want to know more about inheritance. It is now understood that many poor developing countries are characterized by relatively high inequality (World Bank 2005; Ferreira and Ravallion 2009). However, we know far less about how much of that inequality is transmitted across generations. This can create inequalities of opportunity that come with costs to efficiency and growth, as argued in World Bank (2005). Are these economies characterized by a high degree of inter-generational persistence of poverty and affluence, or is there churning associated with successes and failures for adults taking up new economic opportunities? How much do family circumstances matter to adult outcomes? While there has been much debate about these issues in developed countries, they have received relatively little attention in poor, primarily agrarian, economies. Inheritance of agricultural land—the main non-labour factor of production—is the first mechanism one thinks of for the inter-generational transmission of inequality in such settings.

The role of gender is also of interest in this context. We know very little about inter-personal inequality in living standards, including between men and women, given that survey data on consumption are almost invariably collected at the household level. Since adult women are generally married, it is difficult to separate their own welfare from that of their husbands on the basis of household data. However, gender dimensions of inequality are considered important. Male control over land and its inheritance has long been a prominent gender issue in development studies.³ Maternal education and work experience might also be expected to play a role. It has often been argued that

¹ For a dissenting view on the importance of individual land rights *per se* see Platteau (1996). More supportive views can be found in, for example, Feder and Noronha (1987) and Deininger (2003). Also see the discussions in Barrows, and Roth (1990), Besley (1995) and Jacoby and Minten (2007).

² See, for example, Deininger (2003) and Ravallion and van de Walle (2008).

³ In the context of land rights in Africa see Gray and Kevane (1999) for an overview of the issues.

maternal education has an important influence on children's health, nutritional status and schooling.⁴ There might also be implications for adult welfare and economic activity, though there has been less research on this intergenerational linkage.

This study tries to generate new knowledge on the efficiency and equity roles of inheritance in a poor economy. The study takes advantage of an unusual new dataset that measures consumption at a relatively disaggregated level within the household, so that we can build an 'individualized' consumption-based welfare measure, to be matched with individualized data on inheritance and various control variables. The data also allow us to distinguish inheritance and its economic welfare impacts by gender. The data were collected in Senegal—one of many developing countries where these issues have salience.

We document new evidence on the extent of inequality in living standards and its origins, paying special attention to the role of inheritance. Our results lead us to question the traditional model of a developing economy with imperfect credit markets in which privately-owned land is a marketable asset generating an income stream exclusively for the designated individual owner. But our results are easier to understand in the light of richer models of underdeveloped rural economies. We already know from the literature (in anthropology as well as economics) that market failures and non-market allocation processes play an important role in how land is used. As anthropologists have emphasized, agricultural land inheritance in much of Africa is typically filtered through customary land allocation processes involving kinship or community groups.⁵ Inheritance signals a change in responsibility, such that the recipient of inherited land inherits obligations as well as an asset, and it is an asset that is not easily monetized to support other productive investments.

The impacts of inheritance will naturally reflect how the 'dynastic family' allocates its resources, given the market and institutional environment. In principle at least, the extended family has the ability to attain any desired distribution of consumption, independently of the formal assignment of ownership rights. Indeed, it is an open question in this setting whether there is any net impact on the inheritor of a land bequest within the family.

Our results suggest that other mechanisms for the intergenerational transmission of inequality—notably related to parental education and children's schooling—are in fact more important than land inheritance in explaining interpersonal economic welfare and economic activities. Even in very poor settings, parental background can influence the schooling, expectations and life chances generally of children in ways that matter to the realized living standards of adults. There is evidence consistent with this view in the literature. Estudillo et al. (2001) emphasize the combination of both land inheritance and schooling in the inter-generational transmission of wealth in the rural Philippines. Lesorogol et al. (2011) find that the current wealth of Kenyan pastoralists is correlated with parental wealth and formal education but not with livestock inheritance. Ferreira

⁴ See, for example, Haveman and Wolfe (1995), and Hill and King (1995). The causal interpretation of these correlations can be questioned given the possibility of inter-generationally correlated latent factors; see Behrman and Rosenzweig (2002). Dumas and Lambert (2011) find that maternal education plays a weaker role once properly instrumented.

⁵ For an overview of the issues see Shipton and Goheen (1992). Also see the discussion in Platteau (2000).

and Gignoux (2010) find that family background characteristics are an important source of unequal opportunities in Latin America. Dumas and Lambert (2011) find that parental education has a strong effect on child schooling in Senegal. Some degree of intergenerational correlation in occupational choices can be expected, for which we find supportive evidence, though with a (possibly surprising) degree of mobility.

We begin in the following section by discussing the multiple roles of land ownership in this setting. After that we describe our data for Senegal and the methods of analysis, followed by our empirical results.

2. Land inheritance in an underdeveloped rural economy

Land ownership has long played a role in policy-making and thinking about economic development. Taxes and transfers are often conditioned on the amount of land owned. Famously, the American political economist of the late nineteenth century Henry George advocated taxes on the value of land, and (of course) these were to be levied on the designated owner. Taxes on land are found in almost all countries. Also, transfers and various direct interventions are often targeted according to land holding, defined by ownership. These include policies aiming to redistribute land itself. Tenure security is traditionally defined in terms of individual titles of private ownership. There have been many efforts (often supported by external development assistance) to foster individual ownership through land titling, with expected benefits to the government in efforts to tax land value, and also expected gains in both efficiency (promoting land investment through greater tenure security and access to credit) and equity (notably in promoting women's empowerment).⁶

The economic model underlying all these policies essentially views private ownership of land (with or without a formal title) as a form of private wealth, which is expected to deliver exclusively to its owner an income stream derived from the productive capacity of the owned land. In the context of a mainly rural market-based economy, one thus expects land ownership to play an important role in determining the individual's standard of living. Land inheritance is one way of acquiring ownership. Thus land inheritance should be important to the inter-generational transmission of inequality and also to economic activity, including diversification into non-farm production, especially when credit is unavailable. Development policy debates have sometimes focused on inheritance laws, especially reforms aiming to improve women's rights.⁷

However, it is far from clear how relevant this standard model of land as a form of marketable and productive wealth is to underdeveloped and poor rural economies where land markets are thin or non-existent and imperfect (and asymmetric) information and long-established social institutions play an important role in how land is allocated and used. The benefits from efforts to foster individual ownership titles are known to be uncertain when individual titling is introduced in an indigenous system of tenure, which

⁶ On the expected land productivity gains from titling see Feder and Noronha (1987), Barrows and Roth (1990), Besley (1995) and Deininger (2003).

⁷ Hallward-Driemeier and Gajigo (2011), Deininger et al. (2010) and Roy (2011) find evidence that legal reforms related to property rights have brought gains to women (the first in Ethiopia and the other two in India).

is probably why the evidence that such efforts have had their expected benefits appears to be mixed.⁸

Inheritance of the family farm may well bring enhanced individual power within the family—interpretable as a non-pecuniary gain—but it undoubtedly also comes with responsibilities. Past observations about African agriculture lead one to question the extent of the gains to the inheritor of land, who may have to take on various obligations. These naturally include responsibility for the family as an economic unit, but they may also extend well beyond the family. Anthropologists have emphasized the social responsibilities that come with acquired wealth such as through land ownership, notably in Africa.⁹ As Shipton and Goheen (1992: 311) note with reference to land in rural Africa, ‘Rights often entail duties... Cultivation and grazing rights may entail obligations to share farm products beyond the domestic group.’ Similarly, with reference to the Luo people of Kenya, Shipton (1992: 361) argues that ‘Rights of individuals [over land] were not thought sacrosanct, but instead they interlocked with the rights of others, and overlapped with those of families and wider groups.’ Individual responsibilities within a village economy are often embedded in broader social ties, interpretable as means of enforcing co-operative equilibria that bring collective benefits (Platteau 2000). Whether such responsibilities come with a consumption incentive is unclear on *a priori* grounds, given that there are also likely to be non-pecuniary benefits and costs.

The local state and community governance are often involved in land allocation, as are traditional non-market allocation processes. In particular, land that is not kept in use and looked after appropriately risks appropriation by the community in many rural economies: the household head is thus in charge of making sure this doesn’t happen so as to insure the family’s long-term security. Further, membership of a (potentially large) extended family often conveys rights to work the family’s land holding and/or share in its bounty. These arrangements can mean that individual land ownership conveys obligations and associated costs to the owner, especially if he or she is also the head of household. One cannot even rule out consumption losses to the inheritor. Without a land market it will be hard for the recipient to ‘cash in’ the land to finance some other (non-farm) investment. The lack of a land market may then create occupational stickiness, whereby the bequest of land inhibits the recipient’s transition to non-farm activities (though possibly enhancing the scope for such a transition by others in the family). Land-market failures may even entail that the (say) eldest son who gets the land and the responsibilities of being the head of household ends up trapped in farming, while his siblings see new opportunities for diversification into non-farm activities.

Indeed, inheriting the land but without the responsibilities of headship may allow the recipient to leave the land to take up some non-farm activity. Given weak market and governmental institutions for risk-sharing, the family farm is known to serve a social security role in traditional societies. The recipient of the land bequest may then effectively transfer the right to other family members (the mother, spouse and children). Their security (at some minimal level) is thus assured, and the son is free to seek work or start an enterprise elsewhere, such as in an urban area. However, one can also

⁸ Deininger (2003) reviews the evidence. A recent example of a study pointing to success of land titling in raising productivity is Holden et al. (2009); an example finding little or no impact is Jacoby and Minten (2007).

⁹ Though not only Africa; see Ravallion and Dearden (1988), using data for rural Java.

imagine situations in which non-market factors in the allocation of command over the product of land can discourage agriculture, even for the household head, in favour of other (non-farm) activities possibly outside the village economy, such as in urban areas. This can happen when the non-market allocation rules entail a sharing of the product of land, and that the sharing rule entails that inheriting extra land reduces the marginal product of the owner's effort in farming relative to other uses of labour time, thus generating a substitution toward non-farm activities. This is a distortion to intersectoral allocation, in the sense that marginal products of labour become unequal between activities. In principle, such an inefficiency could be avoided if the family is well informed about other (non-farm) income sources, so allowing sharing rules based on total income.

Finally, it is worth underlining that land inheritance can be accompanied by learning within kinship groups—a source of specific human capital that may play an important role in the welfare gains from inheritance. It is widely believed that traditional farming practices in developing countries are characterized by a high degree of farm-specific knowledge, accumulated through experience of farming the same land. Rosenzweig and Wolpin (1985) have emphasized the role of family-specific information in explaining inter-generational and intra-household land transfers, including inheritance; they find support for the claim that specific knowledge about the family farm entails that land is kept within the family rather than being sold. This could also explain why land markets are often thin or non-existent and also why the extended family is so common in underdeveloped rural economies, given that the older generation will have accumulated greater knowledge about the family farm. As long as the extended family can share knowledge there will be little economic loss at the death of the head of the household, though one can imagine circumstances (including unanticipated deaths) when that is not the case.

The upshot of these observations is that inheritance can have ambiguous effects on welfare and economic activities. The rest of this paper will address these issues empirically using an unusual data set for Senegal.

3. Setting and data

More than half of Senegal's population is rural—57 per cent out of a population of about 12.5 million inhabitants in 2009—although the contribution of agriculture to GDP amounts to only 18 per cent.¹⁰ Like other African countries, Senegal has seen considerable population urbanization; at the time of independence in 1960 rural areas were the home of 77 per cent of the population.

On paper, agricultural land is allocated through local administrative processes in Senegal. Since 1964, most of the land (between 95 and 98 per cent) has been part of what is called the national domain (Caverivière 1986). Land use rights are attributed by

¹⁰ This is less than for its neighbouring countries where the share of the rural population is typically about two-thirds (apart for Mauritania which, due to deserts over a large area, has only a third of its population in rural areas) and the share of agriculture in GDP hardly ever below 25 per cent, reaching a third in Mali and even 60 per cent in Guinea-Bissau.

local land committees on the basis of needs and capacity to farm. This land cannot be sold and in theory cannot be bequeathed either (Caverivière 1986).

However, the reality on the ground is clearly rather different. Bequests are common, including land inheritance. The survey data we use indicate that 31 per cent of men, but only 17 per cent of women, report that they inherited land. (If one confines the sample to those whose father has died, the proportions are 43 per cent for men and 28 per cent for women.) Looking at all forms of inheritance (including housing, durables, money and productive assets) 54 per cent of men in the sample inherited something, while 38 per cent of women did. In practice, heirs are given priority to obtain the use rights over any other potential users. Hence, inheritance of paternal lineage land is an important means of access to ownership.

Until the constitution of 2001, women were not allowed to own land in Senegal. Furthermore, until May 2010, they were not allowed to be part of the land committees in charge of the attribution of use rights. Hence, women very rarely received land through this allocation mechanism. On the intergenerational transmission side, several inheritance laws coexist in Senegal that give very different treatment to women. Each individual can choose before his death which law he wants to abide by. The French inspired system of inheritance dictates that wealth be shared equally among children, whatever their gender. By contrast, the Islamic inheritance law (which is the most common choice in practice) limits the inheritance of daughters to half of that of sons.¹¹

In addition, entrenched tradition favours sons for inheriting land. Daughters are supposedly compensated by their brothers with money or other forms of wealth, for what would have otherwise been their share of land inheritance. The data (described in more detail below) reveal that very few women (about 4 per cent) have any land to transfer to their heirs when they die, while more than a third of the men leave some land. In the sample, we observe very few changes in land ownership over the five years preceding the survey. Nevertheless, half of the cases where the amount of land owned increased are due to inheritance.

In these respects, Senegal is not unusual within Sub-Saharan Africa (Cooper 2008, 2010). The rights of women to land are mainly indirect (Platteau et al. 2000). As a daughter living in her father's household, a woman will work on the family land and eventually obtain use rights on a plot. As a wife, she'll work on the land of her husband's family and might also have use rights on a particular plot. If she is in neither of these positions, she simply won't have access to land. This in part explains the high remarriage rate following widowhood or divorce. Whether this fact contributes to gender inequality in terms of well-being is a question we want to explore in this study.

The data used here come from an original survey entitled *Pauvreté et Structure Familiale* (Poverty and Family Structure, henceforth PSF) conducted in Senegal in 2006-07. The PSF survey stems from the co-operation between a team of French

¹¹ Although some of the ethnic groups in Senegal are of matrilineal tradition, those traditions have mostly been displaced by Islam when it comes to inheritance. As a result, inheritance from an adult male other than the father (such as a maternal uncle or foster parent for example) remains a rare occurrence.

researchers and the National Statistical Agency of Senegal.¹² The survey is described in detail in De Vreyer et al. (2008).

The PSF is a nationally representative survey covering a sample of over 1800 households spread over 150 clusters drawn randomly from the census districts so as to insure a geographically representative sample. About 1,750 household records can be exploited, covering 14,450 individuals. The survey describes a population of which the majority (57 per cent) live in rural areas, 48 per cent is male and 95 per cent is Muslim—statistics that accord well with other sources (World Bank 2009). Urbanization over time is evident in that, amongst those adults who had a father in rural areas, 22 per cent now live in urban areas. A similar percentage of those whose mother lived in rural areas also do so.

Senegalese households are large, with slightly more than eight members on average in the PSF. The families are typically multi-generational. Polygamous unions are common, with 24 per cent of married men and 37 per cent of married women engaged in such unions, which mostly comprise a husband and two wives (only 20 per cent of polygamous unions have more than two wives). We find that 31 per cent of polygamous men have non-cohabiting wives. Among those, in only half of the cases are the two parts headed by the same person or by the husband for one and a wife for the other i.e., for the other half, a married polygamous women is living in a household headed by a relative (mainly her father, brother or son). Those large households are extended both horizontally and vertically, with 36 per cent of household members that are neither the head, nor one of his wives or children. Two-thirds of households include such 'extended' family members.

In addition to the usual information on individual characteristics, the survey collected details on each household's structure and budgetary arrangements. Each household was divided into 'cells' according to the following rule: the head of household and unaccompanied dependent members, such as his widowed parent or his children whose mothers do not live in the same household, are grouped together. Then, each wife of the head and her children make up a separate group. Finally, any other family nucleus such as a married child of the household head with his/her spouse and children, forms a separate group. This disaggregation emerged from field interviews as being the relevant way to split the household into its component groups. It is worth noting that enumerators saw this as a fairly natural way to divide households and had no difficulty organizing the household in this way and collecting the data accordingly.

Consumption expenditures are recorded in several parts: first, all common expenditures are collected (housing, electricity, bills etc.). Regarding food expenditures, a detailed account is made of who shares which meal and how much money is specifically used to prepare this meal (the 'DQ', i.e. 'dépenses quotidiennes', which is the name the Senegalese give to the amount of money a woman has at her disposal to buy fresh ingredients for the meals of the day). Then individual consumption is collected at the group level (such as clothing, mobile phone, transportation, food outside the home).

¹² Momar Sylla and Matar Gueye of the Agence Nationale de la Statistique et de la Démographie of Senegal (ANSD) on the one hand and Philippe De Vreyer (University of Paris-Dauphine and IRD-DIAL) Sylvie Lambert (PSE) and Abba Safir (now with the World Bank) designed the survey. The data collection was conducted by the ANSD thanks to the funding of the IDRC (International Development Research Center), INRA Paris and CEPREMAP.

Finally, expenditures that are shared between several groups but not the whole household are collected.

Hence, a measure of per capita consumption can be constructed at the group ('cell') level and this allows us to identify unequal consumption levels within households. Sub-groups also emerge that take some or all of their meals separately (in 17 per cent of households), thus widening the possibility for differences in nutritional intake among household members.

Thanks to these data we can construct a relatively individualized measure of consumption, which is almost never available in household surveys. This is what we will use to assess individual economic welfare. The measure we use here is the amount of expenditures specific to the cell and not shared with any other cell *plus* the cell's imputed share of the household's joint expenditures. We will restrict our study to individuals who are heads of their cells. They are all adults with at least one dependent, and for consumption purposes, they are assumed to be the decision makers at the cell level. We are therefore left with 4401 observations, of which 56.8 per cent are women. The average number of cells per household is 2.51. The range is from 1 to 12; 81 per cent of the sampled households have more than one cell.

When looking at total expenditures, inequalities within the household are evident: the ratio between the expenditures of the richest and the poorest group within a household can be as high as 18 and is still equal to 4.4 after trimming off the 5 per cent most unequal households. Computing an inequality index for the distribution of cash expenditures in the population, we find a Gini index of 59.8 per cent if we attribute to each person the average per capita consumption level in his or her household. The index is 62.7 per cent if instead each individual is attributed the per capita consumption in his cell (i.e. the sum of the per capita expenditures specific to the cell and of the cell's share of common household expenditures, distributed on a per capita basis within the cell). The Gini index of inequality in the distribution of the cell-specific component of cash expenditures (ignoring the joint consumption within the household) is 77.9 per cent.

The individualized consumption data reveal a sizeable gender gap. Regressing the log of cell consumption on gender, the regression coefficient (the difference in mean log consumption) is 0.57 and is significant at the 1 per cent level ($t=14.92$). Importantly for the purpose of our study, these data include information on parental characteristics and inheritance. If the parent has died, the survey asked whether he or she left any inheritance and then, for each person, whether they obtained any inheritance in the form of land, housing, money, durables or productive capital. No valuation of these inheritances was obtained. In particular, we do not know how much land was inherited, although the endogeneity concerns (which we return to) would clearly be even greater using amounts of inherited land rather than simply the incidence of inheritance.

4. Results

We begin by describing inheritance patterns in our data, after which we study the associations with living standards, economic activity and mobility. (Appendix 1 gives summary statistics on the main variables used from the PSF.)

4.1 Inheritance patterns

As can be seen in Table 1, many more fathers have an inheritance to leave to their heirs than do mothers. Seventy-two per cent of deceased fathers left land bequests, but only 22 per cent of deceased mothers did so. Fathers bequeath their wealth to their sons more often than to their daughters, particularly when it is in the form of land. Mothers treated sons and daughters roughly equally in this respect.

With regard to parental characteristics, information was collected on the occupation and education of each parent for each individual. The last place of the mother and father's residence is also known (allowing us to know in particular whether an individual resides in the same village as his parents).

Occupation has been classified under three headings: agriculture, non-agriculture or inactive. Table 2 gives the joint distribution of occupations of parents and their children for men and women. The table also gives Cramer's V; in all four cases the correlation is significant at the 1 per cent level.¹³

We find that 36 per cent of individuals had a father working in agriculture, while 22 per cent declare that their mother was in agriculture. Note that 'housewives' are coded as inactive. It is likely that mothers were declared housewives even when they did some agricultural work so that those who report being in agriculture probably really dedicated most of their time to this occupation, suggesting relatively poor households. When the mother is declared a farmer, in 72 per cent of the cases, the father is also in farming. More generally, parents' sectors of activity are highly correlated.

Women have been moving out of engaging solely in household work. We find that 48 per cent of sampled women were coded as 'inactive' but that this was true of 61 per cent of their mothers. One third of those with inactive mothers went into the non-farm sector, with far fewer going into farming.

In fact Table 2 suggests considerable mobility out of farming across generations. Only one third of the one third of men whose father worked in agriculture stayed in the sector, though there is stronger persistence with respect to the mothers' sector with 43 per cent of the men ('only' 33 per cent of the women) who declared that their mother was working in agriculture doing so as well. Participation in the non-farm sector was more persistent across generations, with nearly three quarters of those men whose father worked in the non-farm sector being also recorded as working in that sector.

To further explore the characteristics of those who inherit, Table 3 gives probits for any form of inheritance, while Table 4 gives probits for land in particular; in both cases these are marginal effects. In each case we also give the breakdown by gender. We give results for two specifications, the second of which drops a number of variables that might be considered endogenous to inheritance. (Later we will use these pruned regressions as the first stage for an instrumental variables estimator.) While the causal interpretation of the first regression (including the endogenous variables) can be questioned, it is still of descriptive interest.

¹³ Cramer's V is a measure of correlation for contingency tables. It lies between 0 and 1 (perfect correlation). The corresponding Chi-square statistic is $2nV^2$ (for our 3x3 table).

We include a wide range of controls in these regressions (and those reported later), including the following individual characteristics: gender, age and age squared, age at first marriage, a dummy for whether one is the first born of a given gender among siblings with the same mother and same father, a dummy for whether one is the first born among all children with the same mother and father, a dummy for whether the first born sibling from the same mother and father is a boy, number of brothers from the same father and mother, number of brothers from the same father only and same mother only, and the same three variables for sisters, dummies for ethnic group, a dummy for being Muslim relative to other religions, having some formal education, whether fostered as a child, and whether fostered at a young age (prior to two years of age, which typically implies a permanent move for the child in the Senegal context). There are also controls for parental characteristics (education, occupation, place of residence, whether the father died in the last two years, and whether the mother did so) and some demographic variables describing the household (log household size) and the individual's cell (log cell size, share of adults and share of children age five and under). In all cases our education variable is defined as a dummy for whether the individual has some formal education.

We continue to find that men are more likely to inherit than women, even with the controls. Being male adds 0.11 to the conditional probability of receiving any inheritance, while it adds 0.07 to the probability of inheriting land.

Unsurprisingly, the death of either parent increases the probability of inheritance, and the coefficients are considerably higher for paternal death. In the full sample, death of the father alone adds 0.67 to the probability of inheritance, while death of the mother adds only 0.15. This is probably due to the fact that at the death of a woman, her land is first passed onto her husband and transferred to the children only at his death. By contrast, children inherit from their father at his death, whether or not their mother is still alive. These effects are significant across almost all strata and specifications, the only exception being that death of the mother is not a significant predictor of land inheritance by women. The effect of a father's recent death dampens the large 'father dead' effect (bringing it down from 0.67 to 0.57, when the mother is still alive), suggesting that inheritance is delayed somewhat.

There is a positive coefficient on education in the regressions for any inheritance, which casts doubt on the idea of substitution by parents between formal schooling and inheritance (whereby some children get some form of inheritance while others get formal schooling). However, there is some sign of such substitution for land inheritance, though it is only statistically significant for women; those women with formal schooling are less likely to inherit land. Obviously, this might reflect the individual choice of an educated woman with a non-farm economic activity to give up her land inheritance to the benefit of her siblings, rather than a parental decision to substitute one form of transmission for another.

Men who were fostered as boys are more likely to inherit land unless they were fostered before aged two. This pattern is plausible. Fostering out a very young child is suggestive of giving away the child (for example to a childless parent), which is an indication that inheritance is unlikely. By contrast, fostering an older child is in general less permanent

and more suggestive of an investment in the child, which would also suggest that inheritance is more likely.¹⁴ None of these effects are statistically significant for girls.

Having a mother active in the non-farm sector significantly increases the probability of any inheritance for men although not for women. For land inheritance, it is father's non-farm activity that matters but negatively—significantly reducing the likelihood of inheriting land for both genders. Paternal activity in farm work has no effect on inheritance but maternal farm work has a positive correlation with men's, and less so women's, land inheritance.

Finally, the number of siblings of the opposite gender is significantly associated with inheriting from one's parents. For men a positive effect on land and on any inheritance, is related to the number of sisters from the same father or same parents;¹⁵ for women, the number of brothers from the same mother reduces the likelihood of getting any inheritance while more brothers from the same father has a significant but small positive influence on land inheritance.

Controlling for these other variables, there is little sign that the probability of inheritance is different between urban and rural areas for men. However, women's probability of any inheritance is lower in rural areas.

4.2 Effects of inheritance on consumption, economic activities and mobility

We regress the log of cell-specific consumption expenditure per person on dummy variables for having inherited land, housing and other assets (finance, consumer durables and physical capital). The regressions also included a large number of control variables, to account for the heterogeneity in individual and household characteristics, including parental characteristics, as described above. Other dependent variables we study include farm and non-farm activity and mobility (whether the respondent lives where his/her parents had lived). We follow past literature in assuming the exogeneity of inheritance (and other aforementioned controls); indeed, inheritance has been used as an instrumental variable for current wealth and land rights in explaining various dimensions of current living standards and land productivity.¹⁶ However, the assumption can be questioned (as we discuss later) and so we also test robustness to relaxing the exogeneity assumption, under assumed exclusion restrictions related to the family of birth.

Tables 5, 6 and 7 give results for (log) cell expenditure per person; recall that this combines the cell-specific expenditures with imputed values for the cell's share of jointly consumed items within the household. Table 5 first presents the coefficients on inheritance estimated without any controls, followed by those estimated by adding various correlates, which we do in two steps: adding controls for geographic effects alone, and then adding the controls for individual, household and parental

¹⁴ See Beck et al. (2011) for detailed discussion of fostering practice in Senegal.

¹⁵ 'Sibling from same parents' means that the siblings share both parents. By contrast, 'sibling from the same father' means that the father is common, but not the mother (and conversely for 'sibling from the same mother').

¹⁶ Examples include Besley (1995) and Akresh et al. (2010).

characteristics. We next give results with all controls for the full sample, and for both an urban-rural stratification and gender stratification in Tables 6 and 7.

Without any controls, inheriting land has a strong negative correlation with consumption (Table 5). But this largely vanishes when one controls for location, reflecting the fact that rural households tend to be both poorer and more likely to inherit land. Similarly, inheriting a house is strongly positively correlated with consumption without controlling for location or other individual and household characteristics, but this vanishes when the controls are added. However, a positive effect of other (non-land, non-housing) forms of inheritance emerges with the controls.

This finding of no statistically significant effects of land or housing inheritance in the full sample with controls is also evident in the various strata (urban/rural, male/female) (Table 6). The significant effects of other types of inheritance in the full sample are confined to the rural stratum and to men.

Recall that we find a sizeable gender gap in consumption. Strikingly, we find that this gender gap in consumption largely vanishes when we add our controls. Adult male heads of cells (typically, though not always, the overall household head as well) have higher consumption than females *ceteris paribus*, but the difference is modest at a gain in log consumption of 0.01, and it is not statistically significant (Table 6). Note, however, that our controls include variables such as schooling, which are unequal between genders. So our finding can be interpreted as indicating that the gender gap in consumption can be explained by the gender difference in individual and household characteristics.

Along with the gender differences in characteristics, which (as we have seen) account for the gender disparity in consumption, there are also gender differences in returns to characteristics. A Chow test rejects the null hypothesis of equality of the coefficients across the regressions for men and women ($F(72,1498) = 1.84$; Prob. $> F < 0.00005$).

Both own education and parental education are significant in the consumption regressions. Strikingly, *maternal* education has a much stronger effect—a change of 0.27 in log consumption—than paternal education, and this is due to its effect in urban areas and for women.

Other parental characteristics matter. Having a father (but not a mother) who worked in the non-farm sector has a large and significant effect on log consumption of 0.30; having a father in farming has a smaller effect (0.17). (The left out category is inactive.) These effects are stronger for men than for women. It is clear that parental characteristics matter, even though land and housing inheritance do not.

Being fostered out as a child is associated with higher adult consumption; the effect is confined to men, and is stronger for those living in urban areas. This result is consistent with the fact that the practice of fostering (which is common in Senegal) is often associated with investment in the human capital of the child. Notice, however, that having been fostered out young has an offsetting effect for men. The positive effect on adult welfare is for those who were not fostered young.¹⁷

¹⁷ On long term impact of fostering in Senegal see Coppoletta et. al (2011)

When we stratify by the interaction of gender with urban-rural residence, a sharper picture emerges on the adult consumption gains from other forms of inheritance, namely that they are confined to rural men (Table 7).

Table 8 summarizes the key marginal effects from probits for agricultural self-employment. Again, parental characteristics matter, though in some possibly surprising ways. Having a father in farming does not have a significant effect, but having a mother who was a farmer makes it more likely one will be a farmer. This is so for men and women, but is significantly larger for men in rural areas. Having had either parent in the non-farm sector makes it less likely one will be a farmer in rural areas, although the effect is far more pronounced for the mother in the case of rural men, while it is the father effect that is stronger for women.

Parental schooling effects on the probability of being a farmer seem weak, though for rural women there is a significant positive effect of father's schooling and negative effect of mother's schooling (both significant at the ten per cent level). Own formal schooling makes it less likely that men will be in farming, but makes this more likely for women.

We find that inheriting land makes farming more likely though this effect is confined to women in rural areas. Endogeneity is a concern here; as women rarely inherit land, those who remained in the same village as their parents and are in farming are more likely to be the ones who inherit land. Inheriting land does not make it more likely that men will be farmers. We find no significant effects of inheriting a house. Other (non-land, non-housing) forms of inheritance are associated with lower probabilities of an adult being in agricultural self-employment. This effect is found in both rural and urban areas, though it is stronger for rural areas (Table 8). The effect is similar in size for men and women, though only statistically significant for women.

Table 9 reports analogous results for non-farm occupations. Having a mother in the non-farm sector significantly enhances the probability of an adult working in the non-farm sector. Higher own schooling increases the probability for men (and it is a larger effect for men in rural areas) but not for women.

We find no evidence that inheriting land has a significant effect on the likelihood of doing non-farm activities (Table 9) (although, as we will see, this changes when we allow land to be endogenous). However, inheritance of other (non-land, non-housing) assets makes it more likely that women will be employed in the non-farm sector, notably in urban areas (Table 11).

A number of effects on mobility—identified by whether an adult lives in the same place as his or her parents—are evident in Table 10. Having a farmer for a father makes a son's mobility more likely, though the effect is only significant for rural men. Having a father who worked in the non-farm sector has the opposite effect—increasing the likelihood of living in the same place, though the effect is only significant for rural areas, and is larger for rural women. Having a father with formal schooling makes mobility more likely for urban men; having had a mother with schooling makes it more likely that a woman will live in the same place as her parents. Own formal schooling

makes mobility more likely, though the effect is only significant for rural men and urban women.

Inheriting land does not have any significant effect on the probability of moving to a location different from where one's father resided in the sample as a whole (Table 10). However, there is a sign that inheriting land actually encourages such mobility for urban women. Inheriting a house makes it more likely a man living in rural areas will have the same residence as his parents—in a rather obvious way since the one brother who will inherit the house is the one who intends to live in it (or already does)—but there are no other significant effects of this form of inheritance. Nor are there any significant effects of other (non-land, non-housing) inheritance.

4.3 Further tests of robustness

We present a number of tests of the robustness of our findings, focusing especially on the effects of inheritance.

Possibly the effect of inheritance is diluted by including in the sample cell heads whose parents are still alive, and cannot (of course) be a source of inheritance. We tested this possibility by only including cases where either the father or the mother is dead. Again we found no significant effect of inheriting land, though a positive effect of inheriting a house did emerge in the urban sample. 'Other inheritance' remained significant in the full sample, and is due to men. Most other results were robust. The effects of maternal characteristics (sector and education) on the probability of being a farmer are stronger when one confines attention to the sample with either parent dead. This sub-sample reveals a stronger effect on non-farm employment of having parents who did non-farm work.

We also tested sensitivity to allowing for an interaction effect between inheritance and the time since the father died; for those with a dead father, the mean time since death is 22 years (the median is 19). It is not clear on *a priori* grounds what one would expect. The inherited asset may have a positive rate of return allowing capital accumulation, though other factors may come into play; for example, there were clearly fewer options to farming for those who inherited the land a long time ago. Also, assets (including land) depreciate in value over time. Our tests involved simply adding an interaction effect between inheritance and years since the death of the father to the preceding regressions.¹⁸ For cell consumption per capita, there was a negative interaction effect though only amongst urban men, for which the effect was significant at the five per cent level. The total effect was positive up to about 30 years, though not significant even when the father died recently. For agricultural self-employment the interaction effect was often positive though generally not significant, except for urban men; for non-farm employment the pattern switched, with a negative interaction effect, though again not strong. However, allowing for such an interaction effect does not change our main findings reported above.

¹⁸ If one assumes that the current value of a past inheritance in amount I is given by $f(t)=[(1+r)(1-d)]^t$ (where r is the rate of return, d is the depreciation rate and t is the number of years since father's death) then the function $f(t)$ can be approximated by a linear function of t with constant parameters if one takes its first-order Taylor series expansion and one assumes that r and d are common across all households. However, these are potentially strong assumptions, especially the constancy of returns.

Yet another variation is to allow for ‘cross-effects’ of inheritance and education of the spouse (for example, including the inheritance of a woman's husband in the regression for her consumption or economic activity). We found no significant cross-effects of inheritance or education on consumption for either men or women. Nor were there any significant effects of a husband's inheritance on the wife's sector of work.¹⁹ There were signs that a wife's inheritance increased the likelihood of men in rural areas being farmers, and made it more likely that urban men would be inactive. One significant cross-effect was that having an educated husband made mobility more likely for women, as measured by whether she lived in the same village as her parents. This was found in both urban and rural areas.

Another possible concern is that some of our regressors may be considered endogenous. The main results on the effects of inheritance were found to be robust to dropping other potentially endogenous variables, namely own-schooling, being fostered as a child and age at first marriage. (The endogeneity concern here is that these variables may be jointly determined with land inheritance.) One change of note is that dropping own education revealed even weaker effects of parental education on sector of employment.

Our assumption that past inheritance is exogenous to current living standards might also be questioned. Choices about who inherits the land may be influenced by factors that are unobserved by us, but observed by the parents or other stakeholders—factors that are also correlated with the economic activity and economic welfare of the child on reaching adulthood. It may be decided by the family group that one of the sons is best suited to taking charge of the family farm on the father's death. This may reflect a latent interest or ability at farming, revealed while growing up. Or it may be that other sons show more aptitude for non-farm work. Parents may also have gender preferences in their choices about inheritance and schooling—choices that are influenced by both market and non-market parameters.²⁰ Another potential source of endogeneity is the fact that children could possibly decline the inheritance, in particular of land, if taking care of it is not compatible with their preferred activity or if, being themselves economically successful, they feel that their siblings have a greater need for it.

In testing the robustness of our results to treating inheritance as endogenous, the key identifying assumption we make is that the death of the father or mother only matters to an individual adult's current economic welfare via inheritance of land or other assets. It is hard to see why parental death sometime in the past would matter to current adult consumption except via inheritance and (hence) wealth. Possibly the shock of parental death will have an impact, but then we control for a father's recent death (within the last two years) in all our regressions.

¹⁹ Among those rural women whose husband inherited land, in most cases the husband also had some other form of (non-land, non-housing) inheritance (there were only 27 exceptions, comprising women whose husband only got some other inheritance). This made it impossible to credibly separate these two forms of spousal inheritance for women, so we aggregated them.

²⁰ For example, using data from the rural Philippines, Estudillo et al. (2001) show how sons are preferred for land inheritance, while daughters are preferred for investments in schooling.

We can only convincingly treat one inheritance variable as endogenous, solving out other endogenous variables.²¹ We do so for any inheritance taken together, and separately for land, excluding other forms of inheritance. We also drop any variables that could be endogenous by the same logic, notably own education. (Parents may decide that one son gets the schooling while the other gets the land.) We also drop fostering and age at marriage for the same reason.

Table 11 gives the IV estimates for each of the dependent variables.²² (Note that the estimator is not feasible for the sub-sample for which the father is dead.) Our results on land inheritance are reasonably robust to relaxing the exogeneity assumption. In particular, we still find that land inheritance does not convey any significant consumption benefit.

However, we now find that land inheritance tends to encourage a shift from farm to non-farm work, suggesting that there was a downward bias in the earlier estimates (Table 9). On investigating this effect further we find that it is present for both the sub-samples that are heads of households and those that are not, but that it was far stronger (in size and statistical significance) for those who are not heads of household. Table 11 gives a split of the results for non-farm activities according to whether or not the cell head is also the overall household head.²³ The difference is even more pronounced if confined to men only; the IV coefficients on land inheritance are 0.460 (s.e.=0.333) for male heads of household as compared to 1.374 (s.e.=0.445) for male non-heads. The latter sub-sample tends to be comprised of married brothers of household heads. When land inheritance comes *without* the responsibilities of headship it appears to be an important factor in encouraging diversification into non-farm activities.

We found that the absence of a significant impact of land inheritance, once instrumented, on the probability of being in farming was true for all the sub-samples we considered (men, women, household heads). Furthermore, on running the same regressions using as the measure of inheritance whether or not the individual received any inheritance of any type, the results were qualitatively very similar to those for land inheritance.

4.4 Implications for explaining inequality

It is clear from these results that bequests can play little role in perpetuating consumption inequality. However, other parental characteristics clearly do matter, both directly and via a child's characteristics at adulthood, notably education. Table 12 presents decompositions of consumption inequality implied by the regressions in Table 6. (Appendix 2 explains how the decompositions were done.)

²¹ We tried using death of father and death of mother as two IVs for two inheritance variables (land and other) but these did not have sufficient power for credible identification.

²² For consumption, we also used the treatment effects model (the 'treatreg' estimator in STATA) which uses full maximum likelihood to estimate the effect of an endogenously chosen binary treatment on another endogenous continuous variable, conditional on two sets of independent variables. This gave very similar results.

²³ No such interaction effects were evident for the other dependent variables in Table 11.

Far more important to inequality than inheritance is ‘own schooling,’ which contributes 9 per cent to overall consumption inequality (almost one fifth of the explained component). The share is even higher for women. Demographics, especially cell size and the proportion of adults, also emerge as large contributors to inequality, especially in rural areas; this probably reflects differences in the available labour force. In the national sample, almost half of the explained inequality is attributable to rural-urban location.

5. Conclusions

We find evidence of significant intergenerational linkages in this setting. This is evident in the correlations between parents’ and children’s sectors of occupation, which persist on adding controls for heterogeneity in other respects. Nonetheless, there is still considerable intergenerational mobility, both across sectors and residentially, which appears to be primarily associated with the transition from farm to non-farm activities. Only one third of the sons of farmers stayed in farming. And adult women in our sample are far more economically active than were their mothers.

Our results suggest that gender plays an important role. There is consumption inequality between men and women, though this is largely accountable to factors such as education. The intergenerational linkages through the mother appear to be stronger than through the father, including on the son’s economic activity. Educated mothers are more likely to have sons in the non-farm sector. While women with formal schooling are no more likely to be in non-farm employment and (slightly) more likely to be farmers, it is their sons who are more likely to find their way into the non-farm sector.

Assuming conditional exogeneity of inheritance (with a wide range of controls), inheriting the land makes it more likely that a woman will stay a farmer, but not so for men. Inheriting other (non-land, non-housing) assets appears to help get women into non-farm work, but there is no such effect for men. However, endogeneity bias might partly account for those results. When we allow for the possible endogeneity of inheritance by assuming that the death of a parent only matters via inheritance (though allowing for the short-term shock of parental death) we find evidence that inheritance does play a role in facilitating diversification into non-farm activities, although this is only present for men who do not also inherit the responsibilities of being the overall head of the household. The potential wealth effect of an inheritance on activity choice seems to be inhibited by the obligations attendant to household headship.

On average, inheriting the land or house brings no significant gain to an adult’s consumption. It appears that intra-household allocation across generations comes fairly close to equalizing consumption between otherwise identical individuals, only one of whom takes on the responsibility for the family’s land and housing assets. However, we find that there are significant gains from inheriting other (non-land and non-housing) assets. In particular, formal schooling appears to yield much higher returns.

In short, while intergenerational linkages clearly matter, there still appears to be considerable intergenerational mobility in this setting. Inheritance of land or housing contributes very little to overall inequality, and does not appear to be an important channel for enhancing economic efficiency through transfers of ownership. Non-land

inheritance, schooling and parental characteristics (especially the mother's) appear to play a far more important role.

Table 1: Inheritance, by gender

<i>Among individual heads of cells whose father or mother is dead</i>	Men	Women	All	t-test of the difference (women-men)
Father has left any form of inheritance	75.44% (983)	68.50% (961)	71.84%	-4.02
Father has transmitted land to this person	43.06% (580)	27.97% (410)	35.19%	-8.48
Mother has left any form of inheritance	21.25% (174)	22.18% (185)	21.72%	0.46
Mother has transmitted land to this person	4.49% (38)	3.35% (29)	3.92%	-1.21

Notes: number of observations in brackets.

Source: authors' computations.

Table 2: Sectoral occupational mobility across generations for men and women

<i>No. observations (% of all individuals)</i>	Men				Women			
	Farm	Non-farm	Inactive	Total	Farm	Non-farm	Inactive	Total
Father's occupation								
Farm	216 (11.37)	313 (16.47)	113 (5.95)	642 (33.79)	228 (9.12)	273 (10.92)	398 (15.91)	899 (35.95)
Non-farm	55 (2.89)	413 (21.74)	96 (5.05)	564 (29.68)	46 (1.84)	397 (15.87)	387 (15.47)	830 (33.19)
Inactive	192 (10.11)	347 (18.26)	155 (8.16)	694 (36.53)	114 (4.56)	231 (9.24)	427 (17.07)	772 (30.87)
Total	463 (24.37)	1,073 (56.47)	364 (19.16)	1,900 (100.00)	388 (15.51)	901 (36.03)	1,212 (48.46)	2,501 (100.00)
Cramer's V				0.178				0.184
Mother's occupation								
Farm	154 (8.11)	133 (7.00)	68 (3.58)	355 (18.68)	185 (7.40)	168 (6.72)	203 (8.12)	556 (22.23)
Non-farm	20 (1.05)	200 (10.53)	40 (2.11)	260 (13.68)	20 (0.80)	215 (8.60)	178 (7.12)	413 (16.51)
Inactive	289 (15.21)	740 (38.95)	256 (13.47)	1,285 (67.63)	183 (7.32)	518 (20.71)	831 (33.23)	1,532 (61.26)
Total	463 (24.37)	1,073 (56.47)	364 (19.16)	1,900 (100.00)	388 (15.51)	901 (36.03)	1,212 (48.46)	2,501 (100.00)
Cramer's V				0.184				0.211

Source: authors' computations.

Table 3: Marginal determinants of the probability of any inheritance

	All		Men		Women	
Male	0.111*** (0.0338)	0.136*** (0.0258)	--	--	--	--
Age	-0.00270 (0.00498)	0.00183 (0.00433)	0.00742 (0.00745)	0.0144** (0.00660)	-0.00201 (0.00614)	0.00148 (0.00550)
Age squared	2.04e-05 (4.70e-05)	-1.71e-05 (4.16e-05)	-7.66e-05 (6.81e-05)	-0.000133** (6.13e-05)	1.22e-05 (6.06e-05)	-2.37e-05 (5.54e-05)
Muslim	0.219*** (0.0530)	0.183*** (0.0510)	0.275*** (0.0960)	0.239** (0.0945)	0.202*** (0.0478)	0.167*** (0.0467)
Serere ethnicity	-0.0806** (0.0406)	-0.0637* (0.0380)	-0.0529 (0.0637)	-0.0447 (0.0609)	-0.114*** (0.0428)	-0.0896** (0.0403)
Poular ethnicity	0.00583 (0.0374)	0.00770 (0.0348)	0.0568 (0.0509)	0.0488 (0.0493)	-0.0461 (0.0432)	-0.0220 (0.0402)
Diola ethnicity	0.0481 (0.0849)	0.00466 (0.0749)	-0.0456 (0.107)	-0.0712 (0.0990)	0.129 (0.103)	0.0680 (0.0882)
Mandingue ethnicity	0.0165 (0.0515)	0.00655 (0.0493)	0.0563 (0.0795)	0.0178 (0.0766)	-0.0284 (0.0537)	-0.0105 (0.0530)
Sarakole ethnicity	0.0408 (0.111)	0.0636 (0.101)	0.0477 (0.137)	0.149 (0.119)	0.0300 (0.144)	0.00827 (0.128)
Mandialque ethnicity	0.187 (0.167)	0.106 (0.157)	0.256 (0.184)	0.231 (0.199)	0.144 (0.189)	0.0475 (0.158)
Other ethnicity	-0.110 (0.0749)	-0.120* (0.0642)	-0.112 (0.123)	-0.117 (0.110)	-0.111 (0.0790)	-0.113* (0.0680)
Brothers same father	-0.000258 (0.00675)	2.14e-05 (0.00645)	-0.00760 (0.0106)	-0.00741 (0.0105)	0.00940 (0.00811)	0.00987 (0.00756)
Brothers same parents	-0.00319 (0.00754)	-0.000256 (0.00710)	-0.0154 (0.0122)	-0.0155 (0.0115)	0.00127 (0.00928)	0.00808 (0.00859)
Sisters same father	0.0235*** (0.00692)	0.0224*** (0.00661)	0.0317*** (0.0111)	0.0359*** (0.0110)	0.0144* (0.00837)	0.0105 (0.00767)
Sisters same parents	0.0132* (0.00744)	0.0142** (0.00697)	0.0268** (0.0110)	0.0267** (0.0106)	0.00634 (0.00909)	0.00555 (0.00854)
Brothers same mother	-0.0337** (0.0148)	-0.0186 (0.0137)	-0.00547 (0.0242)	0.00591 (0.0232)	-0.0523*** (0.0194)	-0.0350** (0.0172)
Sisters same mother	-0.00430 (0.0151)	-0.0183 (0.0140)	-0.0280 (0.0267)	-0.0377 (0.0232)	0.00334 (0.0171)	-0.00792 (0.0167)
First same gender	0.0289 (0.0295)	0.0125 (0.0274)	-0.0272 (0.0542)	-0.0455 (0.0519)	0.0597 (0.0387)	0.0469 (0.0353)
First of siblings	-0.0146 (0.0321)	-0.00173 (0.0303)	-0.0140 (0.0611)	0.0195 (0.0585)	-0.0284 (0.0465)	-0.0378 (0.0426)
First born is male	0.0263 (0.0239)	0.0175 (0.0227)	0.0859* (0.0475)	0.0729 (0.0464)	0.00954 (0.0327)	-0.0168 (0.0294)
Father died recently	-0.111*** (0.0376)	-0.111*** (0.0343)	-0.153** (0.0688)	-0.193*** (0.0597)	-0.0887** (0.0405)	-0.0648* (0.0389)
Father is dead	0.674*** (0.0155)	0.658*** (0.0147)	0.734*** (0.0208)	0.717*** (0.0206)	0.637*** (0.0199)	0.618*** (0.0189)
Mother is dead	0.147*** (0.0279)	0.144*** (0.0264)	0.152*** (0.0424)	0.152*** (0.0404)	0.149*** (0.0346)	0.147*** (0.0324)

Father in farming	-0.0134 (0.0299)	-0.0109 (0.0280)	0.0381 (0.0433)	0.0509 (0.0409)	-0.0479 (0.0354)	-0.0478 (0.0333)
Mother in farm	0.0351 (0.0351)	0.0525 (0.0331)	0.00198 (0.0519)	0.00792 (0.0497)	0.0628 (0.0442)	0.0815** (0.0411)
Father in non-farm	0.0357 (0.0341)	0.0482 (0.0321)	0.00456 (0.0496)	0.0149 (0.0479)	0.0500 (0.0420)	0.0576 (0.0388)
Mother in non-farm	0.0776** (0.0370)	0.0753** (0.0344)	0.140*** (0.0509)	0.143*** (0.0484)	0.0318 (0.0441)	0.0281 (0.0400)
Father's schooling	0.0147 (0.0414)	0.0181 (0.0381)	0.0612 (0.0602)	0.0735 (0.0571)	-0.0103 (0.0486)	-0.00847 (0.0445)
Mother's schooling	-0.0360 (0.0532)	0.00307 (0.0487)	-0.132 (0.0858)	-0.0900 (0.0797)	0.0123 (0.0639)	0.0450 (0.0593)
Father rural	-0.0466 (0.0487)	-0.0574 (0.0447)	-0.114 (0.0752)	-0.142** (0.0706)	-0.00319 (0.0555)	-0.0135 (0.0516)
Mother rural	0.0603 (0.0473)	0.0782* (0.0436)	0.145* (0.0757)	0.161** (0.0721)	0.0163 (0.0541)	0.0319 (0.0498)
Log hh size	0.0333 (0.0229)	0.0187 (0.0204)	0.0562 (0.0344)	0.0249 (0.0311)	0.0146 (0.0265)	0.000344 (0.0240)
Log cell size	0.0347 (0.0236)	0.0370* (0.0218)	0.0597 (0.0394)	0.0591 (0.0369)	0.000396 (0.0334)	0.0168 (0.0300)
Share of cell aged<5	0.0101 (0.0860)	0.0436 (0.0776)	-0.135 (0.167)	0.000953 (0.155)	0.0204 (0.0956)	0.0204 (0.0842)
Share of cell adults	-0.0182 (0.0630)	0.0133 (0.0584)	0.0190 (0.112)	0.0285 (0.106)	-0.0727 (0.0766)	-0.00214 (0.0688)
Has formal schooling	0.0582* (0.0314)	--	0.0869** (0.0441)	--	0.0477 (0.0416)	--
Fostered	0.00747 (0.0424)	--	0.00408 (0.0587)	--	-0.0101 (0.0601)	--
Fostered young	0.0171 (0.0555)	--	0.0722 (0.0909)	--	-0.0224 (0.0679)	--
Age at first marriage	0.00363 (0.00233)	--	0.00441 (0.00326)	--	0.00165 (0.00313)	--
Rural	0.0110 (0.0507)	-0.0316 (0.0475)	0.110 (0.0708)	0.0733 (0.0685)	-0.0803 (0.0565)	-0.110** (0.0514)
No. Observations	3,150	3,524	1,415	1,548	1,735	1,997
Pseudo R ²	0.374	0.373	0.407	0.404	0.381	0.380

Notes: Robust standard errors in parentheses, clustered at the household level; *** p<0.01, ** p<0.05, * p<0.1. The variables 'First same gender', 'First of siblings' and 'First born is male' all refer to children of the same father, same mother. The regression also includes department fixed effects. The reference variables are Wolof ethnicity, all other religions, occupation 'inactive', share of cell members 5-15, no and non-formal schooling.

Source: authors' computations.

Table 4: Marginal determinants of the probability of land inheritance

	All		Men		Women	
Male	0.0748*** (0.0187)	0.0779*** (0.0147)	--	--	--	--
Age	-0.00359 (0.00267)	-0.00248 (0.00215)	-0.000444 (0.00521)	-0.000990 (0.00435)	-0.00249 (0.00225)	-0.00163 (0.00167)
Age squared	3.75e-05 (2.47e-05)	3.06e-05 (2.03e-05)	8.78e-06 (4.66e-05)	1.65e-05 (3.95e-05)	2.43e-05 (2.18e-05)	2.05e-05 (1.67e-05)
Muslim	0.0620*** (0.0206)	0.0506*** (0.0178)	0.0958** (0.0432)	0.0817** (0.0417)	0.0422*** (0.0142)	0.0326*** (0.0115)
Serere ethnicity	0.0310 (0.0249)	0.0189 (0.0207)	0.0889* (0.0508)	0.0651 (0.0447)	0.00322 (0.0189)	-5.75e-05 (0.0142)
Poular ethnicity	0.00156 (0.0184)	-0.00125 (0.0157)	0.0356 (0.0354)	0.0172 (0.0314)	-0.0132 (0.0150)	-0.00503 (0.0122)
Diola ethnicity	0.0585 (0.0625)	0.0365 (0.0487)	0.00777 (0.0832)	0.0124 (0.0748)	0.0913 (0.0768)	0.0434 (0.0488)
Mandingue ethnicity	0.0284 (0.0312)	0.0184 (0.0262)	0.0576 (0.0607)	0.0196 (0.0507)	0.00548 (0.0230)	0.00777 (0.0193)
Sarakole ethnicity	-0.0133 (0.0475)	-0.0123 (0.0373)	-0.0120 (0.0893)	0.0121 (0.0768)	-0.0113 (0.0338)	-0.0113 (0.0247)
Mandiaque ethnicity	0.250 (0.229)	0.204 (0.200)	0.442 (0.303)	0.445 (0.317)	0.150 (0.176)	0.0799 (0.118)
Other ethnicity	-0.0405 (0.0334)	-0.0292 (0.0282)	-0.00740 (0.0785)	-0.0181 (0.0635)	-0.0414*** (0.0154)	-0.0276** (0.0137)
Brothers same father	0.00630* (0.00333)	0.00476 (0.00292)	0.00747 (0.00675)	0.00470 (0.00623)	0.00655** (0.00293)	0.00473** (0.00239)
Brothers same parents	-0.000305 (0.00408)	-0.00111 (0.00346)	-0.00412 (0.00844)	-0.00522 (0.00745)	-0.00114 (0.00364)	-0.000291 (0.00283)
Sisters same father	0.00386 (0.00346)	0.00523* (0.00298)	0.0120* (0.00693)	0.0155** (0.00625)	-0.000384 (0.00300)	0.000449 (0.00237)
Sisters same parents	0.00669* (0.00386)	0.00586* (0.00324)	0.0178** (0.00743)	0.0169** (0.00668)	0.000874 (0.00349)	9.87e-05 (0.00277)
Brothers same mother	-0.0149** (0.00741)	-0.0160** (0.00642)	0.00604 (0.0164)	-0.00185 (0.0154)	-0.0269*** (0.00821)	-0.0224*** (0.00674)
Sisters same mother	-0.00548 (0.00845)	-0.00466 (0.00694)	-0.0208 (0.0172)	-0.0128 (0.0148)	-0.00119 (0.00685)	-0.000725 (0.00544)
First same gender	0.0168 (0.0149)	0.00978 (0.0127)	0.00467 (0.0356)	0.000149 (0.0321)	0.0190 (0.0148)	0.0114 (0.0118)
First of siblings	-0.00495 (0.0164)	-0.000204 (0.0144)	-0.0170 (0.0394)	-0.0103 (0.0361)	0.00292 (0.0175)	0.00665 (0.0146)
First born is male	0.0154 (0.0125)	0.0112 (0.0106)	0.0531* (0.0289)	0.0395 (0.0269)	0.00628 (0.0123)	0.00392 (0.00958)
Father died recently	-0.0397** (0.0161)	-0.0346** (0.0135)	-0.0922*** (0.0303)	-0.0802*** (0.0263)	-0.0193 (0.0127)	-0.0141 (0.0102)
Father is dead	0.334*** (0.0141)	0.317*** (0.0133)	0.429*** (0.0196)	0.412*** (0.0190)	0.259*** (0.0179)	0.234*** (0.0163)
Mother is dead	0.0364** (0.0158)	0.0257* (0.0135)	0.0674** (0.0289)	0.0558** (0.0262)	0.0177 (0.0140)	0.00852 (0.0106)
Father in farming	-0.0122 (0.0149)	-0.00769 (0.0128)	0.00233 (0.0283)	0.0128 (0.0253)	-0.0151 (0.0130)	-0.0118 (0.0104)
Mother in farm.	0.0465** (0.0201)	0.0471*** (0.0176)	0.0953** (0.0411)	0.0810** (0.0372)	0.0211 (0.0169)	0.0244* (0.0141)

Father in non-farm	-0.0577*** (0.0167)	-0.0514*** (0.0144)	-0.0845*** (0.0310)	-0.0746*** (0.0278)	-0.0383*** (0.0144)	-0.0330*** (0.0117)
Mother in non-farm	0.000288 (0.0206)	0.00290 (0.0177)	0.0154 (0.0424)	0.0147 (0.0382)	-0.00356 (0.0165)	-0.00374 (0.0130)
Father's schooling	-0.00155 (0.0245)	-0.00489 (0.0202)	-0.0590 (0.0413)	-0.0667* (0.0351)	0.0213 (0.0255)	0.0156 (0.0201)
Mother's schooling	-0.00649 (0.0312)	-0.0138 (0.0236)	-0.0434 (0.0537)	-0.0282 (0.0481)	0.00812 (0.0335)	-0.0126 (0.0186)
Father rural	0.0398* (0.0227)	0.0283 (0.0193)	0.0158 (0.0483)	-0.000487 (0.0436)	0.0392** (0.0184)	0.0289* (0.0149)
Mother rural	0.0486** (0.0228)	0.0543*** (0.0193)	0.0865* (0.0466)	0.0876** (0.0416)	0.0216 (0.0192)	0.0270* (0.0153)
Log hh size	0.0347*** (0.0121)	0.0209** (0.00967)	0.0655*** (0.0230)	0.0385** (0.0195)	0.0156 (0.0103)	0.00842 (0.00777)
Log cell size	0.00386 (0.0115)	0.00511 (0.00951)	-0.00968 (0.0251)	-0.0122 (0.0223)	-0.00406 (0.0116)	0.000672 (0.00878)
Share of cell aged<5	-0.0625 (0.0451)	-0.0535 (0.0378)	-0.274** (0.131)	-0.241** (0.116)	-0.0228 (0.0330)	-0.0143 (0.0253)
Share of cell adults	-0.0299 (0.0326)	-0.0221 (0.0276)	-0.161** (0.0768)	-0.148** (0.0692)	-0.0135 (0.0264)	-0.00708 (0.0205)
Has formal schooling	-0.0213 (0.0145)	--	-0.0119 (0.0286)	--	-0.0232* (0.0119)	--
Fostered	0.0489** (0.0244)	--	0.119** (0.0488)	--	-0.0211 (0.0174)	--
Fostered young	-0.0537*** (0.0200)	--	-0.108*** (0.0344)	--	-0.00254 (0.0257)	--
Age at first marriage	0.000873 (0.00111)	--	0.00210 (0.00195)	--	-0.000342 (0.00125)	--
Rural	0.0114 (0.0247)	0.00836 (0.0209)	0.0769* (0.0452)	0.0748* (0.0401)	-0.0170 (0.0201)	-0.0141 (0.0154)
No. observations	3,150	3,524	1,415	1,548	1,699	1,976
Pseudo R ²	0.378	0.363	0.389	0.388	0.359	0.367

Notes: Robust standard errors in parentheses, clustered at the household level; *** p<0.01, ** p<0.05, * p<0.1. The variables 'First same gender', 'First of siblings' and 'First born is male' all refer to children of the same father, same mother. Regressions also contain regional (department) dummies. The reference variables are Wolof ethnicity, all other religions, occupation 'inactive', share of cell members 5-15, no and non-formal schooling.

Source: authors' computations

Table 5: Estimated effects of inheritance on log cell per capita consumption with and without controls

	(1) No controls	(2) Rural location and department dummies	(3) As in (2) + controls for individual and household characteristics
Inherited land	-0.482*** (0.0702)	-0.0620 (0.0572)	-0.0655 (0.0585)
Inherited house	0.327*** (0.0606)	0.0955* (0.0488)	0.0515 (0.0498)
Other inheritance	0.0631 (0.0656)	0.167*** (0.0521)	0.106** (0.0526)
Constant	11.45*** (0.0381)	12.28*** (0.303)	11.91*** (0.383)
Observations	4,326	4,326	3,558
R ²	0.016	0.376	0.499

Source: authors' computations.

Table 6: Regressions for log cell expenditure per capita

	(1) Full sample	(2) Rural	(3) Urban	(4) Men	(5) Women
Male	0.0137 (0.0562)	0.000143 (0.0770)	0.0711 (0.0850)	--	--
Age	0.00472 (0.00963)	0.0107 (0.0127)	0.00586 (0.0140)	0.0264 (0.0161)	-0.0274** (0.0116)
Age squared	-0.000111 (9.89e-05)	-0.000217 (0.000132)	-6.26e-05 (0.000139)	-0.000352** (0.000155)	0.000289** (0.000121)
Muslim	0.223* (0.124)	-0.154 (0.224)	0.343** (0.137)	0.304* (0.184)	0.165 (0.127)
Serere ethnicity	-0.350*** (0.0827)	-0.610*** (0.147)	-0.156* (0.0914)	-0.461*** (0.109)	-0.243** (0.0968)
Poular ethnicity	-0.114 (0.0699)	-0.222** (0.113)	-0.0258 (0.0860)	-0.213** (0.0957)	-0.0288 (0.0786)
Diola ethnicity	-0.265* (0.156)	-0.567 (0.534)	-0.0682 (0.152)	-0.118 (0.221)	-0.366** (0.151)
Mandingue ethnicity	-0.125 (0.0966)	-0.0702 (0.159)	-0.164 (0.117)	-0.156 (0.134)	-0.0884 (0.105)
Sarakole ethnicity	-0.00777 (0.167)	0.0638 (0.299)	0.0186 (0.228)	-0.0864 (0.230)	0.0374 (0.187)
Mandiaque ethnicity	-0.362* (0.190)	-0.551*** (0.163)	-0.190 (0.250)	-0.0677 (0.285)	-0.459** (0.225)
Other ethnicity	0.0434 (0.142)	0.0361 (0.200)	-0.0897 (0.145)	0.220 (0.189)	-0.0854 (0.143)
Brothers same father	0.00307 (0.0109)	-0.00771 (0.0148)	0.00710 (0.0160)	0.00506 (0.0182)	-0.000408 (0.0135)
Brothers same parents	0.0157 (0.0134)	0.0176 (0.0190)	0.00880 (0.0190)	0.0312 (0.0207)	0.00431 (0.0171)
Sisters same father	-0.00188 (0.0111)	-0.00907 (0.0163)	0.000743 (0.0153)	0.00145 (0.0195)	0.00182 (0.0132)
Sisters same parents	0.0286** (0.0126)	0.0365* (0.0209)	0.0276* (0.0164)	0.0183 (0.0200)	0.0379** (0.0167)
Brothers same mother	0.0215 (0.0244)	-0.0416 (0.0428)	0.0573* (0.0297)	0.0181 (0.0380)	0.0270 (0.0312)
Sisters same mother	0.00623 (0.0292)	0.0337 (0.0424)	-0.00222 (0.0404)	-0.0210 (0.0451)	0.0117 (0.0361)
First same gender	0.00601 (0.0468)	0.0437 (0.0667)	-0.0336 (0.0660)	0.0467 (0.0882)	-0.0108 (0.0657)
First of siblings	-0.0228 (0.0520)	-0.117* (0.0704)	0.0842 (0.0788)	-0.0870 (0.0964)	0.0397 (0.0809)
First born is male	0.0545 (0.0417)	0.0895 (0.0570)	0.0153 (0.0583)	0.0723 (0.0865)	0.0898 (0.0560)
Father died recently	-0.0578 (0.0850)	-0.0145 (0.128)	-0.0761 (0.112)	-0.111 (0.165)	-0.0853 (0.0920)
Father in farming	0.170*** (0.0574)	0.153** (0.0722)	0.192** (0.0952)	0.215*** (0.0822)	0.108 (0.0692)
Mother in farm	-0.0145 (0.0600)	0.0552 (0.0769)	-0.109 (0.0952)	-0.168* (0.0875)	0.108 (0.0731)
Father in non-farm	0.295*** (0.0593)	0.315*** (0.0962)	0.259*** (0.0776)	0.386*** (0.0862)	0.217*** (0.0728)
Mother in non-farm	-0.0823 (0.0619)	0.131 (0.100)	-0.171** (0.0797)	-0.119 (0.0994)	-0.0608 (0.0727)

Father's schooling	0.0367 (0.0702)	-0.249** (0.124)	0.110 (0.0813)	0.0925 (0.121)	0.00213 (0.0818)
Mother's schooling	0.265*** (0.0933)	0.205 (0.184)	0.371*** (0.113)	0.182 (0.144)	0.303** (0.125)
Father rural	0.0958 (0.0725)	0.102 (0.114)	0.0347 (0.0937)	0.130 (0.138)	0.0718 (0.0830)
Mother rural	-0.129* (0.0699)	-0.108 (0.107)	-0.119 (0.0945)	-0.104 (0.134)	-0.143* (0.0808)
Log hh size	-0.154*** (0.0430)	-0.0743 (0.0646)	-0.245*** (0.0563)	-0.129** (0.0561)	-0.167*** (0.0518)
Log cell size	-0.246*** (0.0398)	-0.417*** (0.0620)	-0.164*** (0.0523)	-0.234*** (0.0688)	-0.212*** (0.0590)
Share of cell aged<5	-0.563*** (0.141)	-0.579*** (0.196)	-0.440** (0.205)	-0.784*** (0.301)	-0.561*** (0.166)
Share of cell adults	0.696*** (0.116)	0.371** (0.168)	0.824*** (0.162)	0.627*** (0.214)	0.671*** (0.139)
Has formal schooling	0.400*** (0.0497)	0.336*** (0.0843)	0.396*** (0.0620)	0.315*** (0.0761)	0.468*** (0.0662)
Fostered	0.187** (0.0751)	0.123 (0.109)	0.199* (0.103)	0.305*** (0.110)	0.0413 (0.101)
Fostered young	-0.0334 (0.0966)	0.00367 (0.142)	-0.0545 (0.130)	-0.298** (0.150)	0.177 (0.130)
Age at first marriage	0.00422 (0.00395)	0.0118** (0.00590)	-0.00149 (0.00523)	-0.00420 (0.00534)	0.0173*** (0.00537)
Inherited land	-0.0527 (0.0609)	0.0410 (0.0858)	-0.0959 (0.0896)	-0.0825 (0.0910)	0.00956 (0.0762)
Inherited house	0.0164 (0.0519)	-0.104 (0.0836)	0.0962 (0.0666)	0.00775 (0.0765)	0.0160 (0.0650)
Other inheritance	0.126** (0.0547)	0.137** (0.0674)	0.0914 (0.0860)	0.237*** (0.0758)	0.0189 (0.0685)
Rural	-0.670*** (0.0896)			-0.669*** (0.125)	-0.722*** (0.100)
Constant	11.68*** (0.379)	10.55*** (0.727)	10.63*** (0.539)	11.26*** (0.627)	12.15*** (0.461)
Observations	3,165	1,637	1,528	1,419	1,746
R ²	0.504	0.301	0.394	0.479	0.529

Notes: Robust standard errors in parentheses, and are clustered at the household level. *** p<0.01, ** p<0.05, * p<0.1. The variables 'First same gender', 'First of siblings' and 'First born is male' all refer to children of the same father, same mother. Regressions also contain regional (department) dummies. The reference variables are Wolof ethnicity, all other religions, occupation 'inactive', share of cell members 5-15, no and non-formal schooling.

Source: authors' computations

Table 7: Regressions for log cell expenditure per capita by gender and rural/urban residence

	(1) Rural men	(2) Rural women	(3) Urban men	(4) Urban women
Age	0.0239 (0.0232)	-0.0170 (0.0143)	0.0344 (0.0212)	-0.0247 (0.0181)
Age squared	-0.000393* (0.000222)	0.000163 (0.000148)	-0.000343* (0.000203)	0.000300 (0.000185)
Muslim	-0.265 (0.290)	-0.106 (0.233)	0.473** (0.194)	0.245 (0.163)
Serere ethnicity	-0.822*** (0.189)	-0.458*** (0.164)	-0.249* (0.135)	-0.0306 (0.118)
Poular ethnicity	-0.353** (0.149)	-0.129 (0.133)	-0.0922 (0.126)	0.0166 (0.0994)
Diola ethnicity	-0.532 (0.547)	-0.525 (0.673)	0.159 (0.237)	-0.219 (0.146)
Mandingue ethnicity	-0.0685 (0.200)	-0.0484 (0.181)	-0.217 (0.182)	-0.137 (0.129)
Sarakole ethnicity	-0.0465 (0.516)	0.332 (0.463)	0.0841 (0.309)	-0.00152 (0.218)
Mandjaque ethnicity	-0.553 (0.382)	-0.593*** (0.180)	0.336 (0.342)	-0.363 (0.314)
Other ethnicity	0.0997 (0.253)	-0.0493 (0.225)	0.136 (0.215)	-0.255 (0.157)
Brothers same father	-0.00759 (0.0263)	-0.00231 (0.0182)	0.0186 (0.0258)	-0.00920 (0.0209)
Brothers same parents	0.0302 (0.0315)	0.00264 (0.0241)	0.0300 (0.0287)	-0.00723 (0.0252)
Sisters same father	-0.0278 (0.0295)	0.00545 (0.0189)	0.0128 (0.0256)	0.00100 (0.0196)
Sisters same parents	0.0287 (0.0314)	0.0516* (0.0275)	0.00979 (0.0266)	0.0414* (0.0223)
Brothers same mother	-0.0556 (0.0764)	-0.0386 (0.0528)	0.0858** (0.0402)	0.0382 (0.0406)
Sisters same mother	0.0814 (0.0954)	0.00493 (0.0372)	-0.0801* (0.0440)	0.0323 (0.0569)
First same gender	0.0862 (0.120)	0.0477 (0.0960)	0.0910 (0.121)	-0.0959 (0.0937)
First of siblings	-0.191 (0.135)	-0.0484 (0.107)	-0.0309 (0.135)	0.160 (0.127)
First born is male	0.183 (0.118)	0.114 (0.0739)	0.00841 (0.120)	0.0584 (0.0870)
Father died recently	-0.137 (0.250)	-0.00607 (0.157)	-0.0534 (0.227)	-0.149 (0.113)
Father in farming	0.114 (0.107)	0.170** (0.0852)	0.407*** (0.131)	-0.0131 (0.126)
Mother in farm	-0.0397 (0.114)	0.147 (0.0899)	-0.399*** (0.148)	0.113 (0.130)
Father in non-farm	0.443*** (0.134)	0.196* (0.118)	0.384*** (0.109)	0.148 (0.102)
Mother in non-farm	-0.0221 (0.143)	0.291** (0.140)	-0.158 (0.130)	-0.185** (0.0917)
Father's schooling	-0.386	-0.0897	0.225	0.0521

	(0.246)	(0.149)	(0.140)	(0.0971)
Mother's schooling	0.270	0.0745	0.255	0.425***
	(0.279)	(0.227)	(0.183)	(0.153)
Father rural	0.329	0.0238	-0.0932	0.0946
	(0.262)	(0.122)	(0.164)	(0.111)
Mother rural	-0.399	-0.00812	0.0434	-0.212*
	(0.259)	(0.118)	(0.160)	(0.116)
Log hh size	0.0667	-0.174**	-0.309***	-0.176**
	(0.0890)	(0.0752)	(0.0719)	(0.0756)
Log cell size	-0.458***	-0.349***	-0.104	-0.202**
	(0.108)	(0.0880)	(0.0918)	(0.0862)
Share of cell aged<5	-1.079**	-0.484**	-0.293	-0.580**
	(0.425)	(0.232)	(0.455)	(0.241)
Share of cell adults	0.246	0.334	0.815***	0.716***
	(0.307)	(0.203)	(0.311)	(0.205)
Has formal schooling	0.189	0.446***	0.327***	0.446***
	(0.124)	(0.114)	(0.101)	(0.0827)
Fostered	0.270*	-0.0535	0.350**	0.0617
	(0.151)	(0.174)	(0.169)	(0.125)
Fostered young	-0.301	0.228	-0.293	0.172
	(0.221)	(0.203)	(0.218)	(0.166)
Age at first marriage	0.00190	0.0273***	-0.0114	0.0111
	(0.00881)	(0.00760)	(0.00697)	(0.00799)
Inherited land	0.00869	0.113	-0.156	-0.0340
	(0.121)	(0.113)	(0.147)	(0.111)
Inherited house	-0.165	-0.0896	0.150	0.0746
	(0.111)	(0.110)	(0.110)	(0.0822)
Other inheritance	0.259***	0.0293	0.176	-0.0160
	(0.0982)	(0.0796)	(0.122)	(0.110)
Constant	11.31***	11.65***	9.912***	11.94***
	(1.124)	(0.627)	(0.733)	(0.563)
Observations	733	904	686	842
R ²	0.297	0.321	0.385	0.395

Notes: Robust standard errors in parentheses, clustered at the household level. *** p<0.01, ** p<0.05, * p<0.1. The variables 'First same gender', 'First of siblings' and 'First born is male' all refer to children of the same father, same mother. Regressions also contain regional (department) dummies. The reference variables are Wolof ethnicity, all other religions, occupation 'inactive', share of cell members 5-15, no and non-formal schooling.

Source: authors' computations

Table 8: Marginal determinants of agricultural employment including inheritance, schooling and parental characteristics

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
	All	Rural	Urban	Men	Women	Rural men	Rural women	Urban men	Urban women
Father in farming	0.0181 (0.0167)	0.0219 (0.0326)	0.00746 (0.0102)	0.0120 (0.0262)	0.0188 (0.0166)	0.00892 (0.0494)	0.0334 (0.0403)	0.0158 (0.0197)	-0.00185 (0.00154)
Mother in farm	0.0845*** (0.0242)	0.145*** (0.0397)	0.00958 (0.0172)	0.139*** (0.0376)	0.0446** (0.0218)	0.240*** (0.0570)	0.0779* (0.0432)	-0.0226* (0.0118)	0.0204 (0.0186)
Father in non-farm	-0.0355* (0.0197)	-0.0799* (0.0468)	-0.00620 (0.0103)	-0.0471 (0.0310)	-0.0438** (0.0207)	-0.150** (0.0687)	-0.0516 (0.0618)	0.00564 (0.0171)	-0.0119 (0.00737)
Mother in non-farm	-0.0414* (0.0224)	-0.107* (0.0557)	-0.00494 (0.00936)	-0.0656* (0.0356)	-0.0226 (0.0219)	-0.161* (0.0838)	-0.0878 (0.0625)	-0.0138 (0.0136)	-0.000187 (0.00154)
Father's schooling	0.0371 (0.0305)	0.150* (0.0855)	-0.00126 (0.0103)	0.0260 (0.0514)	0.0282 (0.0313)	0.0285 (0.131)	0.211** (0.107)	0.00739 (0.0223)	-0.000617 (0.00149)
Mother's schooling	-0.0186 (0.0396)	-0.104 (0.0803)	-0.000787 (0.0163)	-0.0228 (0.0618)	-0.0171 (0.0357)	-0.127 (0.117)	-0.136* (0.0725)	-0.0194 (0.0187)	0.0376 (0.0343)
Own schooling	-0.0114 (0.0187)	-0.0452 (0.0455)	-0.00598 (0.00835)	-0.0504* (0.0282)	0.0435* (0.0242)	-0.0543 (0.0669)	0.0230 (0.0620)	-0.0604*** (0.0193)	0.00654 (0.00492)
Inherited land	0.0623*** (0.0230)	0.0916** (0.0426)	0.0117 (0.0137)	0.0201 (0.0304)	0.0992*** (0.0331)	0.00354 (0.0585)	0.168*** (0.0653)	0.00841 (0.0198)	0.0115 (0.00955)
Inherited house	0.00736 (0.0188)	0.0175 (0.0388)	0.0137 (0.0124)	-0.0111 (0.0265)	0.0209 (0.0209)	-0.0305 (0.0533)	0.0842 (0.0548)	0.0120 (0.0152)	-0.000296 (0.00207)
Inherited other	-0.0432*** (0.0159)	-0.0760** (0.0351)	-0.0149** (0.00754)	-0.0331 (0.0250)	-0.0289* (0.0155)	-0.0596 (0.0496)	-0.0573 (0.0435)	-0.0149 (0.0139)	-0.00198 (0.00156)
Observations	3,190	1,652	1,433	1,422	1,702	731	866	627	554
Pseudo R ²	0.305	0.213	0.234	0.305	0.365	0.218	0.284	0.298	0.479

Notes: Robust standard errors in parentheses, clustered at the household level; *** p<0.01, ** p<0.05, * p<0.1. Marginal effects are reported. The regressions include controls listed in Table 5.

Source: authors' computations

Table 9: Marginal determinants of non-agricultural employment including inheritance, schooling and parental characteristics

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
	All	Rural	Urban	Men	Women	Rural men	Rural women	Urban men	Urban women
Father in farming	0.0239 (0.0274)	0.00973 (0.0295)	0.0474 (0.0389)	0.0496 (0.0378)	0.00176 (0.0347)	0.0369 (0.0468)	-0.0150 (0.0334)	0.0414 (0.0422)	0.0710 (0.0625)
Mother in farm	-0.0199 (0.0318)	-0.0329 (0.0326)	0.0443 (0.0482)	-0.0801* (0.0445)	0.0272 (0.0402)	-0.114** (0.0504)	0.0125 (0.0377)	0.0300 (0.0523)	0.0420 (0.0728)
Father in non-farm	0.0486* (0.0294)	0.128*** (0.0460)	0.00437 (0.0343)	0.102** (0.0436)	0.0237 (0.0388)	0.202*** (0.0749)	0.0994 (0.0616)	0.0353 (0.0395)	9.27e-06 (0.0522)
Mother in non-farm	0.135*** (0.0313)	0.130** (0.0545)	0.125*** (0.0302)	0.176*** (0.0476)	0.113*** (0.0394)	0.240*** (0.0889)	0.0623 (0.0623)	0.105*** (0.0317)	0.131*** (0.0470)
Father's schooling	-0.0479 (0.0330)	-0.0960* (0.0540)	0.0114 (0.0356)	-0.00532 (0.0612)	-0.0591 (0.0393)	0.0161 (0.124)	-0.0980* (0.0555)	0.0329 (0.0453)	-0.00552 (0.0510)
Mother's schooling	0.0115 (0.0513)	0.173* (0.0946)	-0.0323 (0.0559)	0.0843 (0.0854)	-0.0237 (0.0582)	0.268** (0.128)	0.135 (0.121)	0.0321 (0.0692)	-0.0619 (0.0769)
Own schooling	0.0483* (0.0269)	0.0814* (0.0452)	0.0276 (0.0297)	0.161*** (0.0380)	-0.0130 (0.0339)	0.247*** (0.0747)	0.0266 (0.0508)	0.117*** (0.0351)	-0.0383 (0.0432)
Inherited land	-0.0121 (0.0317)	0.00565 (0.0364)	0.0185 (0.0444)	0.0154 (0.0440)	-0.00516 (0.0419)	0.0735 (0.0582)	-0.0372 (0.0426)	-0.0216 (0.0487)	0.0726 (0.0632)
Inherited house	0.0150 (0.0281)	-0.0101 (0.0352)	0.00849 (0.0345)	-0.00725 (0.0396)	0.0239 (0.0351)	-0.0129 (0.0551)	0.00961 (0.0437)	-0.0205 (0.0383)	0.0250 (0.0477)
Inherited other	0.0392 (0.0279)	-0.00184 (0.0313)	0.104*** (0.0350)	0.0189 (0.0372)	0.0783** (0.0367)	0.00184 (0.0482)	0.0269 (0.0390)	0.0631* (0.0349)	0.152*** (0.0536)
Observations	3,169	1,624	1,538	1,409	1,747	705	893	682	849
Pseudo R ²	0.208	0.142	0.154	0.259	0.179	0.217	0.145	0.149	0.148

Notes: Robust standard errors in parentheses, clustered at the household level; *** p<0.01, ** p<0.05, * p<0.1. Marginal effects are reported. The regressions include controls listed in Table 5.

Source: authors' computations

Table 10: Marginal determinants of living in the same residence including inheritance, schooling and parental characteristics

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
	All	Rural	Urban	Men	Women	Rural men	Rural women	Urban men	Urban women
Father in farming	-0.0765*** (0.0259)	-0.0607* (0.0319)	-0.0367 (0.0533)	-0.0884** (0.0357)	-0.0523 (0.0351)	-0.0427** (0.0208)	-0.0411 (0.0475)	-0.0346 (0.0776)	-0.0508 (0.0695)
Mother in farm	-0.0296 (0.0297)	-0.0187 (0.0347)	-0.0266 (0.0700)	0.000695 (0.0396)	-0.0466 (0.0384)	0.0190 (0.0190)	-0.0587 (0.0488)	-0.0442 (0.101)	-0.0130 (0.0924)
Father in non-farm	0.0747** (0.0299)	0.116*** (0.0423)	0.0700* (0.0418)	0.0545 (0.0380)	0.0761* (0.0398)	0.0547*** (0.0164)	0.200*** (0.0710)	0.0681 (0.0652)	0.0597 (0.0530)
Mother in non-farm	0.0987*** (0.0308)	0.148*** (0.0433)	0.0887** (0.0417)	0.0999** (0.0409)	0.0605 (0.0392)	0.0487*** (0.0179)	0.149* (0.0820)	0.139* (0.0725)	0.0735 (0.0514)
Father's schooling	-0.0258 (0.0370)	0.0924 (0.0662)	-0.0823** (0.0406)	-0.0599 (0.0519)	-0.0248 (0.0461)	0.0157 (0.0479)	0.0634 (0.119)	-0.131** (0.0624)	-0.0511 (0.0505)
Mother's schooling	0.0472 (0.0483)	0.138** (0.0658)	-0.0385 (0.0566)	0.0462 (0.0720)	-0.00292 (0.0629)	0.0425* (0.0226)	0.211* (0.123)	-0.0343 (0.122)	-0.0865 (0.0644)
Own schooling	-0.0482* (0.0277)	-0.0512 (0.0509)	-0.0732** (0.0354)	-0.0636* (0.0378)	-0.0222 (0.0366)	-0.0783* (0.0456)	0.00202 (0.0692)	-0.0916 (0.0590)	-0.0908** (0.0455)
Inherited land	-0.0380 (0.0290)	0.00648 (0.0392)	-0.102** (0.0444)	0.00195 (0.0389)	-0.0901** (0.0402)	0.0161 (0.0236)	0.0239 (0.0631)	-0.0337 (0.0728)	-0.142** (0.0573)
Inherited house	0.0310 (0.0251)	-0.00573 (0.0366)	0.0246 (0.0365)	0.0475 (0.0326)	0.00455 (0.0343)	0.0418* (0.0220)	-0.0932 (0.0584)	0.0149 (0.0587)	0.0202 (0.0460)
Inherited other	0.0167 (0.0261)	0.0297 (0.0340)	0.00577 (0.0442)	0.0221 (0.0343)	-0.0217 (0.0357)	0.0259 (0.0211)	0.0150 (0.0536)	0.0143 (0.0695)	-0.00543 (0.0571)
Observations	3,190	1,652	1,538	1,382	1,761	702	912	663	849
Pseudo R ²	0.173	0.287	0.330	0.326	0.109	0.360	0.202	0.415	0.314

Notes: robust standard errors in parentheses, clustered at the household level; *** p<0.01, ** p<0.05, * p<0.1. Marginal effects are reported. The regressions include controls listed in Table 5.

Source: authors' computations.

Table 11: IV estimates for all dependent variables, land inheritance

	(1) Log expenditure per capita	(2) Non-farm All	(3) Heads	(4) Non-heads	(5) Farm employment	(6) Same residence as parents
Inherited land	0.0417 (0.145)	0.142** (0.0592)	0.101 (0.0905)	0.142* (0.0778)	-0.0368 (0.0437)	-0.0325 (0.0598)
Father in farming	0.142*** (0.0546)	0.0260 (0.0218)	-0.00304 (0.0301)	0.0536* (0.0275)	0.0163 (0.0195)	-0.0656*** (0.0207)
Mother in farm.	-0.0147 (0.0581)	-0.0231 (0.0236)	-0.0374 (0.0333)	-0.0203 (0.0308)	0.117*** (0.0247)	-0.0120 (0.0237)
Father in non-farm	0.331*** (0.0586)	0.0583** (0.0238)	0.0456 (0.0347)	0.0782** (0.0315)	-0.0111 (0.0168)	0.0514** (0.0246)
Mother in non-farm	-0.0865 (0.0599)	0.109*** (0.0249)	0.165*** (0.0399)	0.0982*** (0.0313)	-0.0182 (0.0154)	0.0960*** (0.0255)
Father's schooling	0.173** (0.0683)	-0.0182 (0.0260)	0.0217 (0.0391)	-0.0476 (0.0364)	0.0179 (0.0158)	-0.00635 (0.0299)
Mother's schooling	0.275*** (0.0858)	-0.0138 (0.0380)	-0.00311 (0.0603)	-0.0103 (0.0485)	0.0151 (0.0253)	0.0447 (0.0366)
Constant	12.07*** (0.361)	0.0324 (0.108)	0.350 (0.224)	0.000805 (0.135)	0.541*** (0.117)	0.598*** (0.180)
Observations	3,499	3,524	1,519	2,005	3,524	3,524
R ²	0.494	0.234	0.278	0.270	0.272	0.214

Notes: Linear IV coefficients for (2)-(6). Robust standard errors in parentheses; *** p<0.01, ** p<0.05, * p<0.1. The regressions exclude own education, whether fostered and age at first marriage; other controls are as in Table 5.

Source: authors' computations.

Table 12: Inequality decomposition implied by Table 6

	Share of inequality attributable to each source (%):				
	All	Rural	Urban	Men	Women
Male	0.17	0.00	1.56	--	--
Age	0.82	0.90	1.82	-1.50	-5.67
Age squared	-1.45	-1.09	-1.74	4.14	5.54
Muslim	-0.06	-0.21	0.32	-0.09	-0.08
Serere	1.39	9.16	-0.06	1.97	0.97
Poular	0.54	0.50	0.16	1.91	0.05
Diola	-0.06	1.41	-0.03	-0.04	-0.15
Mandingue	0.11	-0.04	0.41	0.13	0.06
Sarakole	0.00	0.18	0.02	-0.05	0.02
Mandiaque	-0.07	0.32	0.05	-0.04	-0.01
Other ethnicity	0.02	0.07	-0.01	0.31	0.03
Brothers same father	0.06	-0.05	0.27	0.16	0.00
Brothers same parents	0.14	0.35	0.06	0.66	0.01
Sisters same father	-0.04	-0.08	0.03	0.04	0.03
Sisters same parents	0.36	0.69	0.26	0.24	0.70
Brothers same mother	0.13	0.02	0.43	0.07	0.24
Sisters same mother	0.03	-0.02	0.00	-0.07	0.09
First same gender	0.00	-0.18	-0.08	-0.07	-0.01
First of siblings	-0.04	0.51	0.53	0.00	0.06
First born is male	0.21	1.00	0.13	0.05	-0.09
Father dead recently	0.02	0.03	0.10	0.00	0.01
Father in farming	-1.82	-0.21	-1.03	-1.70	-1.46
Mother in farming	0.16	-0.40	0.67	1.82	-1.21
Father in non-farm	5.75	3.22	4.21	7.46	4.80
Mother in non-farm	-0.40	0.79	0.38	-0.51	-0.40
Father's schooling	0.34	-0.25	1.20	0.87	0.02
Mother's schooling	1.07	0.38	2.43	0.44	1.85
Father rural	-2.53	-0.22	-0.27	-3.42	-2.10
Mother rural	3.46	0.30	0.87	2.71	4.33
Log hh size	3.00	1.11	8.46	3.07	2.80
Log cell size	6.39	27.57	7.62	4.85	3.44
Share of cell aged<5	3.62	7.14	4.21	2.93	3.20
Share of cell adults	10.63	11.15	18.11	5.69	9.02
Has formal schooling	9.21	3.75	10.15	7.04	10.88
Fostered	0.91	0.72	1.14	1.41	0.22
Fostered young	-0.06	0.00	-0.08	-0.42	0.55
Age at first marriage	0.99	4.28	-0.42	-0.47	2.64
Inherited land	0.25	0.31	0.28	0.90	-0.03
Inherited house	0.07	-0.76	0.41	-0.01	0.10
Other inheritance	0.08	1.44	0.36	-0.49	0.01
Rural	23.80	--	na	23.46	27.57
Total share explained	50.40	30.10	39.40	47.90	52.90

Note: All sources do not add to the total share explained due to the omission of the share of inequality due to department of residence.

Source: authors' computations.

References

- Akresh, Richard, Joyce J. Chen and Charity Moore (2010). 'Altruism, Co-operation and Intrahousehold Allocation: Agricultural Production in Polygynous Households'. mimeo, University of Illinois, Urbana-Champaign.
- Atkinson, Anthony B. (1970). 'On the Measurement of Inequality'. *Journal of Economic Theory*, 2: 244-263.
- Barrows, R. and M. Roth (1990). 'Land Tenure and Investment in African Agriculture: Theory and Evidence'. *Journal of Modern African Studies*, 28(2): 265 - 297.
- Beck, Simon, De Vreyer Philippe, Lambert Sylvie, Marazyan Karine and Abla Safir, (2011). 'Child Fostering in Senegal'. mimeo, Paris School of Economics.
- Behrman, Jere and Mark Rosenzweig (2002). 'Does Increasing Women's Schooling Raise the Schooling of the Next Generation?' *American Economic Review*, 92(1): 323-334.
- Besley, Timothy (1995). 'Property Rights and Investment Incentives: Theory and Evidence from Ghana'. *Journal of Political Economy*, 103(5): 903-937.
- Caverivière, Monique (1986). 'Incertitudes et Devenir du Droit Foncier Sénégalais'. *Revue Internationale de Droit Comparé*, 38(1): 95-115.
- Cooper, Elisabeth (2008). 'Inheritance Practices and the Intergenerational Transmission of Poverty: A Literature Review and Annotated Bibliography'. London: Overseas Development Institute (ODI) and Chronic Poverty Research Center (CPRC).
- Cooper, Elisabeth (2010). 'Inheritance and the Intergenerational Transmission of Poverty in Sub-Saharan Africa: Policy Considerations'. Chronic Poverty Research Center (CPRC).
- Coppoletta Rosalinda, De Vreyer Philippe, Lambert Sylvie and Abla Safir (2011). 'The Long Term Impact of Child Fostering in Senegal: Adults Fostered in their Childhood'. mimeo, Paris School of Economics.
- Deininger, Klaus (2003). *Land Policies for Growth and Poverty Reduction*, Washington, DC: World Bank.
- Deininger, Klaus, Aparajita Goyal and Hari Nagarajan (2010). 'Inheritance Law Reform and Women's Access to Capital'. Policy Research Working Paper 5338, Washington, DC: World Bank.
- DeVreyer, Philippe, Sylvie Lambert, Abla Safir and Momar B. Sylla (2008). 'Pauvreté et Structure Familiale, Pourquoi une Nouvelle Enquête?' *Stateco*, 102: 261-275.
- Dumas, Christelle and Sylvie Lambert (2011). 'Children's Education in Senegal: How Does Family Background Influence Achievement?' *Journal of African Economies*, 20(1): 1-26.
- Estudillo, Jonna P., Agnes R. Quisumbing and Keijiro Otsuka (2001). 'Gender Differences in Land Inheritance and Schooling Investments in the Rural Philippines'. *Land Economics*, 77(1): 130-143.
- Feder, Gershon and Raymond Noronha (1987). 'Land Rights Systems and Agricultural Development in Sub-Saharan Africa'. *World Bank Research Observer*, 2 (2): 143-169.

- Ferreira, Francisco and Jérémie Gignoux (2010). 'The Measurement of Inequality of Opportunity: Theory and an Application to Latin America'. Policy Research Working Paper 4659, Washington, DC: World Bank.
- Ferreira, Francisco and Martin Ravallion (2009). 'Poverty and Inequality: The Global Context'. In Wiemer Salverda, Brian Nolan and Tim Smeeding (eds.), *The Oxford Handbook of Economic Inequality*, Oxford: Oxford University Press.
- Fields, Gary (1996). 'Accounting for Differences in Income Inequality'. mimeo, Cornell University.
- Gray, Leslie and Michael Kevane (1999). 'Diminished Access, Diverted Exclusion: Women and Land Tenure in Sub-Saharan Africa'. *African Studies Review*, 42(2): 15-39.
- Hallward-Driemeier, Mary, and Ousman Gajigo (2011). 'Strengthening Economic Rights and Women's Occupational Choice: The Impact of Reforming Ethiopia's Family Law'. Paper Presented at 25th Conference of Center for the Study of African Economies, University of Oxford.
- Haveman, Robert and Barbara Wolfe (1995). 'The Determinants of Children's Attainments: A Review of Methods and Findings'. *Journal of Economic Literature*, 33(4): 1829-78.
- Hill, Anne M. and Elizabeth King (1995). 'Women's Education and Economic Well-Being'. *Feminist Economics*, 1(2): 21-46.
- Holden, Stein T., Klaus Deininger and Hosaena Hagos Ghebru (2009). 'Impacts of Low-Cost Land Certification on Investment and Productivity'. *American Journal of Agricultural Economics*, 91(2): 359-373.
- Jacoby, Hanan and Bart Minten (2007). 'Is Land Titling in Sub-Saharan Africa Cost-Effective? Evidence from Madagascar'. *World Bank Economic Review*, 21 (3): 461-485.
- Lesorogol, Carolyn, Gina Chowa and David Ansong (2011). 'Livestock or the Pen: The Effects of Inheritance and Education on Poverty Among Pastoralists'. Working Paper 188, Chronic Poverty Research Centre, University of Manchester.
- Platteau, Jean-Philippe (1996). 'The Evolutionary Theory of Land Rights as Applied to Sub-Saharan Africa: A Critical Assessment'. *Development and Change*, 27(1): 29-86.
- Platteau, Jean-Philippe (2000). *Institutions, Social Norms and Economic Development*, Amsterdam: Harwood Academic Publishers.
- Platteau, Jean-Philippe, Abraham, A., Braselle, A., Gaspart, F., Niang, A., Sawadogo J. and L. Stevens (2000). 'Marriage System, Access to Land, and Social Protection for Women, Part One: The Case of Senegal'. Namur University, Belgium, Centre de Recherche en Economie du Développement.
- Ravallion, Martin and Shaohua Chen (1999). 'When Economic Reform is Faster than Statistical Reform: Measuring and Explaining Income Inequality in Rural China'. *Oxford Bulletin of Economics and Statistics*, 61(1): 33-56.
- Ravallion, Martin and Lorraine Dearden (1988) 'Social Security in a 'Moral Economy': An Empirical Analysis for Java'. *The Review of Economics and Statistics*, 70: 36-45.

- Ravallion, Martin and Dominique van de Walle (2008) *Land in Transition: Reform and Poverty in Rural Vietnam*: Palgrave Macmillan.
- Rosenzweig, Mark and Kenneth Wolpin (1985). 'Specific Experience, Household Structure and Intergenerational Transfers: Farm Family Land and Labour Arrangements in Developing Countries'. *Quarterly Journal of Economics*, 100: 961-87.
- Roy, Sanchari (2011). 'Empowering Women: Inheritance Rights and Female Education in India'. Paper Presented at 25th Conference of Centre for the Study of African Economies, University of Oxford.
- Shipton, Parker (1992). 'Debts and Trespasses: Land, Mortgages, and the Ancestors in Western Kenya'. *Africa: Journal of the International African Institute*, 62(3): 357-388.
- Shipton, Parker and Mitzi Goheen (1992). 'Understanding African Land-Holding: Power, Wealth, and Meaning'. *Africa: Journal of the International African Institute* 62(3): 307-325.
- World Bank (2005). *World Development Report 2006: Equity and Development*. Washington, DC: World Bank.
- World Bank (2009). *World Development Indicators*, Washington, DC: World Bank.

Appendix Table 1: Descriptive statistics

Variable	Obs.	Mean	Std. Dev.	Min	Max
Log cell consumption per capita	4326	11.326	1.466	5.336	18.059
In farming	4377	0.222	0.416	0	1
In non-farm	4377	0.413	0.492	0	1
Same residence as parents	4377	0.555	0.497	0	1
Male	4377	0.437	0.496	0	1
Age	4370	42.314	15.063	13	96
Muslim	4377	0.954	0.210	0	1
Serere ethnicity	4364	0.128	0.334	0	1
Poular ethnicity	4364	0.308	0.462	0	1
Diola ethnicity	4364	0.040	0.197	0	1
Mandingue ethnicity	4364	0.062	0.241	0	1
Sarakole ethnicity	4364	0.023	0.149	0	1
Mandiaque ethnicity	4364	0.011	0.104	0	1
Other ethnicity	4364	0.040	0.196	0	1
Brothers same father (no.)	4322	1.668	2.316	0	21
Brothers same parents (no.)	4329	1.986	1.633	0	11
Sisters same father (no.)	4317	1.502	2.216	0	17
Sisters same parents (no.)	4330	1.952	1.671	0	14
Brothers same mother (no.)	4321	0.318	0.892	0	10
Sisters same mother (no.)	4321	0.277	0.826	0	10
First born same gender	4377	0.465	0.499	0	1
First born of siblings	4377	0.291	0.454	0	1
First born is male	4280	0.569	0.495	0	1
Father dead recently	4377	0.068	0.252	0	1
Father in farming	4377	0.051	0.220	0	1
Mother in farming	4377	0.374	0.484	0	1
Father in non-farm	4377	0.221	0.415	0	1
Mother in non-farm	4377	0.282	0.450	0	1
Father's schooling	4377	0.136	0.343	0	1
Mother's schooling	4377	0.102	0.302	0	1
Father rural	3877	0.051	0.220	0	1
Mother rural	3948	0.660	0.474	0	1
Log hh size	4377	0.651	0.477	0	1
Log cell size	4377	2.142	0.689	0	3.784
Share of cell members aged<5	4377	0.970	0.699	0	2.708
Share of adults in cell	4377	0.146	0.213	0	0.8
Has formal education	4377	0.645	0.318	0	1
Fostered	4377	0.262	0.440	0	1
Fostered young	4377	0.152	0.359	0	1
Age at first marriage	3838	0.073	0.261	0	1
Inherited land	4377	22.394	6.773	10	56
Inherited house	4377	0.246	0.431	0	1
Other inheritance	4377	0.334	0.472	0	1
Rural	4377	0.073	0.261	0	1
Father dead	4274	0.574	0.494	0	1
Mother dead	4266	0.657	0.475	0	1

Notes: The statistics are population weighted. The variables 'First born same gender', 'First born of siblings' and 'First born is male' all refer to children of the same father, same mother.

Appendix Table 2: Descriptive statistics by gender

Variable	Women					Men					t-test of equality
	Obs	Mean	Std.	Min	Max	Obs	Mean	Std.	Min	Max	
Log cell consumption p.c.	2442	11.093	1.370	6.812	17.589	1884	11.625	1.529	5.336	18.059	-14.47
In farming	2471	0.170	0.376	0	1	1906	0.289	0.453	0	1	-8.10
In non farm	2471	0.331	0.471	0	1	1906	0.519	0.500	0	1	-14.16
Same residence	2471	0.465	0.499	0	1	1906	0.671	0.470	0	1	-11.66
Age	2471	39.224	14.124	13	96	1902	46.291	15.302	15	95	-18.73
Muslim	2471	0.951	0.215	0	1	1906	0.957	0.203	0	1	-0.88
Serere ethnicity	2466	0.129	0.335	0	1	1898	0.126	0.332	0	1	0.40
Poular ethnicity	2466	0.295	0.456	0	1	1898	0.325	0.468	0	1	-2.16
Diola ethnicity	2466	0.045	0.207	0	1	1898	0.034	0.182	0	1	2.13
Mandingue ethnicity	2466	0.069	0.253	0	1	1898	0.053	0.225	0	1	2.64
Sarakole ethnicity	2466	0.023	0.150	0	1	1898	0.023	0.149	0	1	-0.02
Mandiaque ethnicity	2466	0.013	0.112	0	1	1898	0.009	0.092	0	1	1.73
Other ethnicity	2466	0.039	0.193	0	1	1898	0.042	0.200	0	1	-0.81
Brothers same father	2440	1.628	2.226	0	21	1882	1.719	2.428	0	20	-1.51
Brothers same parents	2444	2.045	1.645	0	10	1885	1.911	1.614	0	11	1.90
Sisters same father	2435	1.505	2.198	0	17	1882	1.498	2.239	0	15	-0.12
Sisters same parents	2445	2.038	1.676	0	12	1885	1.842	1.658	0	14	4.00
Brothers same mother	2440	0.326	0.891	0	10	1881	0.308	0.894	0	9	1.14
Sisters same mother	2440	0.282	0.843	0	10	1881	0.270	0.804	0	10	1.05
First same gender	2471	0.448	0.497	0	1	1906	0.488	0.500	0	1	-2.34
First of siblings	2471	0.261	0.439	0	1	1906	0.329	0.470	0	1	-4.80
First born is male	2411	0.433	0.496	0	1	1869	0.743	0.437	0	1	-20.34
Father died recently	2471	0.074	0.262	0	1	1906	0.061	0.240	0	1	1.52
Mother died recently		0.054	0.227	0	1		0.047	0.211	0	1	1.42
Father in farming	2471	0.384	0.486	0	1	1900	0.360	0.480	0	1	1.64
Mother in farm.	2471	0.238	0.426	0	1	1900	0.200	0.400	0	1	3.33
Father in non-farm	2471	0.299	0.458	0	1	1900	0.261	0.439	0	1	2.92

Mother in non-farm	2471	0.146	0.354	0	1	1900	0.122	0.328	0	1	2.74
Father's schooling	2471	0.116	0.320	0	1	1900	0.083	0.276	0	1	3.35
Mother's schooling	2471	0.057	0.233	0	1	1900	0.043	0.202	0	1	2.17
Father rural	2191	0.644	0.479	0	1	1683	0.681	0.466	0	1	-1.98
Mother rural	2236	0.632	0.482	0	1	1709	0.676	0.468	0	1	-2.43
Log hh size	2471	2.179	0.630	0	3.784	1906	2.094	0.756	0	3.784	5.21
Log cell size	2471	1.223	0.601	0	2.708	1906	0.645	0.681	0	2.708	28.01
Share of cell aged<5	2471	0.209	0.233	0	0.8	1906	0.065	0.149	0	0.75	22.71
Share of adults in cell	2471	0.502	0.287	0	1	1906	0.829	0.254	0.143	1	-36.00
Has formal education	2471	0.229	0.420	0	1	1900	0.305	0.461	0	1	-7.08
Fostered	2471	0.143	0.350	0	1	1900	0.164	0.371	0	1	-1.68
Fostered young	2471	0.088	0.284	0	1	1900	0.054	0.226	0	1	4.29
Age at first marriage	2138	18.684	4.788	10	51	1699	26.997	6.007	12	56	-48.66
Inherited land	2471	0.176	0.381	0	1	1906	0.336	0.472	0	1	-11.29
Inherited house	2471	0.278	0.448	0	1	1906	0.407	0.491	0	1	-8.45
Inherited other	2471	0.049	0.215	0	1	1906	0.105	0.307	0	1	-6.76
Rural	2471	0.576	0.494	0	1	1906	0.573	0.495	0	1	1.36
Father is dead	2412	0.600	0.490	0	1	1862	0.730	0.444	0	1	-8.67
Mother is dead	2414	0.359	0.480	0	1	1852	0.462	0.499	0	1	-6.83

Notes: The statistics are population weighted. The variables 'First born same gender', 'First born of siblings' and 'First born is male' all refer to children of the same father, same mother.

Appendix 2: Regression-based decomposition of inequality

A measure of (relative) inequality can be written in the generic form:

$$I = I(y_1 / \mu, \dots, y_N / \mu)$$

where y_i is the i 'th person's consumption in a population of size N and μ is mean income. We assume that this measure is continuous, symmetric (swapping incomes does not change the measure), normalized such that inequality is zero when all persons have the same income, and that the measure satisfies the Pigou-Dalton transfer axiom' such that a transfer from rich to poor reduces inequality. For some sorts of distributional comparisons we may not need to know any more about the measure of inequality.²⁴

In our empirical work we will focus on two special cases of the above class of measures. The first is the well-known Gini index (G), given by the (household-size weighted) mean absolute deviation between all pairs of per capita household incomes. The second is a member of the Generalized Entropy class of additively decomposable measures, namely the average log deviation of incomes from their mean:

$$LD = \frac{1}{N} \sum_{i=1}^N \ln(\mu / y_i)$$

We will ask how much of the level of inequality or its change over time is due to some variable determining consumption through a stochastic process. Write the regression model as:

$$y_i = \sum_{k=1}^m \beta_k x_{ik}$$

where x_{ik} is the k 'th predictors (x_{im} can be taken to be an error term, with $\beta_m=1$). Following Fields (1996) and Ravallion and Chen (1999), the contribution of the k 'th explanatory variable to total inequality is given by:

$$c_k = \frac{\beta_k \text{cov}(x_k, y)}{\text{var}(y)}$$

This is simply the product of the partial regression coefficient of income on schooling (holding all other variables constant) with that total regression coefficient of schooling on income (holding nothing else constant). The contributions of each asset to the changes over time can then be determined. The precise decomposition will naturally depend on the regression specification.

²⁴ For example, if the Lorenz curve (giving, on the vertical axis, the share of total income held by the poorest $x\%$ of the population) for distribution A is everywhere above that of B then all inequality measures in the above class of measures will show higher inequality in B than A (Atkinson 1970).